

"TESTING COMPETITIVE MARKET STRUCTURES"

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## ABSTRACT

An accurate understanding of the structure of competition is important in the formulation of many marketing strategies. For example, in new product launch, product reformulation, or positioning decisions, the strategist wants to know which of his competitors will be most affected and hence most likely to respond. Many marketing science models have been proposed to identify market structure.

In this paper we examine the managerial problem and propose a criterion by which to judge an identified market structure. Basically, our criterion is a quantification of the intuitive managerial criterion that a "submarket" is a useful conceptualization if it identifies which products are most likely to be affected by "our" marketing strategies. We formalize this criterion within the structure of classical hypothesis testing so that a marketing scientist can use statistical statements to evaluate a market structure identified by: (1) behavioral hypotheses, (2) managerial intuition, or (3) market structure identification algorithms.

Mathematically, our criterion is based on probabilities of switching to products in the situation where an individual's most preferred product is not available. 'Submarkets' are said to exist when consumers are statistically more likely to buy again in that 'submarket' than would be predicted based on an aggregate "constant ratio" model. For example, product attributes (e.g., brand, form, size), use situations (e.g., coffee in the morning versus coffee at dinner), and user characteristics (e.g., heavy versus light users) are specified as hypotheses for testing alternate competitive structures.

Measurement and estimation procedures are described and a convergent approach is illustrated. An application of the methodology to the coffee market is presented and managerial implications of six other applications are described briefly.

## PERSPECTIVE

The modeling of competitive market structure represents an interesting area for research on market behavior and a crucial activity in the formulation of effective, marketing strategy. New product development, product policy, and competitive advertising and pricing decisions depend in part upon the identification of which products compete most strongly with one another.

Recently, two lines of research have addressed the structure of competition among products. The first is based on information processing theory and mathematical psychology. Bettman (1971, 1979), Haines (1974), Lussier and Olshavsky (1979), and Payne (1976) describe the decision processes that individual consumers use to attain information, assimilate that information, and utilize it to make product decisions. Such decision processes are often described by sequential processing hierarchies. In mathematical psychology, Tversky (1972) has developed a theory called "elimination-by-aspects" (EBA) in which product characteristics are chosen at random and all products not having those characteristics are eliminated. The process continues until one product remains. While EBA sometimes looks like a decision hierarchy and has often been cited as a basis for aggregate market structure, Tversky and Sattath (1979, page 540) point out that the process is not entered sequentially, but rather at random. Tversky and Sattath (1979) instead propose a sequential processing rule, called "hierarchical elimination model" (HEM), which applies to very special cases called preference trees. Hauser and Tversky (1983) extend HEM to the general case. However, neither HEM, nor EBA is preserved by aggregation. Each needs further assumptions to be applicable to an aggregate market structure.

The second line of research does not model explicitly the individual decision sequence, but rather describes the aggregate nature of competition. Various approaches to defining criteria for competition have been proposed including, among others; (1) switching between brands (Butler 1976; Kalwani and Morrison 1977; Rao and Sabavala 1978; Rubinson, Vanhonacker and Bass 1979; Charnes, Cooper, Learner and Phillips 1979;

Vanhonacker 1979, 1980; Ehrenberg and Goodhart 1982)); (2) the structure of choice probabilities (McFadden 1980, Batsell 1980); (3) in-use substitution (Stefflre 1972; Day, Shocker and Srivastava 1979; Bourgeois, Haines, and Sommers 1979; Srivastava, Leone, and Shocker 1981; Arabie, Carroll, DeSarbo, and Wind 1981); (4) the segments of consumers who use the product (Frank, Massy, and Wind 1972); and (5) similarity of interpurchase times (Fraser and Bradford 1983). These approaches allow competitive structures to be defined by either product attributes (e.g., form or brand), use similarity, or consumer characteristics (e.g., heavy versus light users).

The information processing/mathematical psychology literature has often been cited as a behavioral motivation for models of aggregate market structure. Although the aggregate descriptions may be consistent with individual behavior, this is rarely true. For example, Tversky and Sattath (1979, p.552) point out, "It is well known that most probabilistic models (including EBA and the constant ratio rule) are not preserved by aggregation." Even at the individual level, Hauser and Tversky (1983) caution "switching hypotheses are neither implied by nor imply cognitive processing hierarchies." The individual cognitive structure and aggregate market structure methods reflect two different approaches to understanding the bounds of competition.<sup>1</sup>

Determining the correct aggregate competitive structure is important to managers. Consider a product line decision. It is often desirable for a firm to have one product in each of the major sectors of the market, and to avoid unnecessary duplication between products. If a firm can identify a sector in which it does not now compete, then it could consider allocating resources to develop a new product for introduction in that sector in order to generate incremental sales and profits. Conversely, the presence of

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<sup>1</sup>Stochastic process models such as Jeuland, Bass, and Wright (1980) are based on assumptions about the aggregate summary probability distribution (e.g. Dirchlet) of individual behavior. Nonetheless, aggregate parameters (e.g., the parameters of the Dirchlet) are estimated and the market is described by those aggregate parameters. We classify this method as an aggregate market structure method.

duplication within a sector could lead to the dropping of a product and a consolidation of the product line within that sector. Changes in the product line reflect commitment of millions of dollars in many organizations. Similarly, the success of advertising, pricing or product reformulation strategies in a competitive environment depends in part upon which competitors are most (least) affected by the strategy and hence most (least) likely to consider a competitive response.

Managerial actions will depend on the specific market structure. It is important to have a procedure that can determine if all products in a market compete with each other or if 'submarkets' exist where the level of competition is high within them and low between them<sup>2</sup>. It is equally important to know how the submarkets are identified if they exist. Are the submarkets characterized by product attributes (decaffeinated versus caffeinated coffee), user characteristics (heavy versus light coffee consumption rates) or uses (morning coffees versus evening coffees)? Finally, it is important that the manager have confidence that the market structure upon which he (she) plans his (her) strategy is a reasonable description of the probable actions by consumers who are affected by his (her) strategy.

This paper pursues the research thrust directed at testing the aggregate competitive structure of a market once hypotheses about structure have been generated. It begins with a managerially relevant definition of market structure and formulates the definition as a mathematical statement based on switching when a product is deleted from the market. We develop a statistical test to identify whether a market satisfies that mathematical definition. We use the statistical test in a procedure for testing aggregate competitive structures. We describe alternative measurement procedures to obtain the necessary data for the tests and provide illustrative examples of testing

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<sup>2</sup>Testing a 'market' for 'submarkets' presupposes that a 'market' has already been defined. For example, at one level we can test the market, 'coffees,' for submarkets. At another level 'coffees' can be a submarket of 'hot beverages' or even 'beverages.' If the appropriate data are available our tests can be used to address the latter question as well as the former.

competitive structures based on attribute, use, and user characteristics within the product deletion definitional framework.

We outline managerial use of the statistics and report an empirical application based on laboratory measures for the coffee market. We illustrate empirically how our methodology can test alternative theories of market structure. In this section, we also examine whether our methodology converges with some selected alternative methodologies. Finally, we relate our experience by briefly summarizing the managerial implications of six other applications of our procedures. This paper closes with a brief discussion of research issues.

#### DEFINITION OF COMPETITIVE STRUCTURE

As an intuitive introduction to the issues of competitive structure consider the following automobile industry situation. In 1981, most of the sporty front wheel drive (FWD) automobiles were sold in the U.S. by foreign manufacturers, e.g., Honda, Saab, Volkswagon. In June 1981, General Motors launched the FWD J-body cars (Chevrolet Cavalier and Pontiac J-2000) which were similar in appearance to the sporty FWD foreign cars. However, the J-body cars were similar in price, interior room, power, and, of course, manufacturer to General Motors' family FWD X-body cars (Chevrolet Citation, Pontiac Phoenix). If the J-body cars drew customers from sporty FWD imported cars, then General Motors would have been successfully competing against the imported cars. If, instead, the J-body cars drew customers from family FWD domestic cars, then the J-body cars would have cannibalized sales from the X-body cars. Figure 1 is a visual representation of the alternative placements of the J-body cars.<sup>3</sup> (Figure 1 is simplified for expositional purposes. There are over 160 automobile models on the market.)

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<sup>3</sup> Figure 1 assumes that the structure is stable in the sense that the introduction of the new car does not totally redefine other groupings.

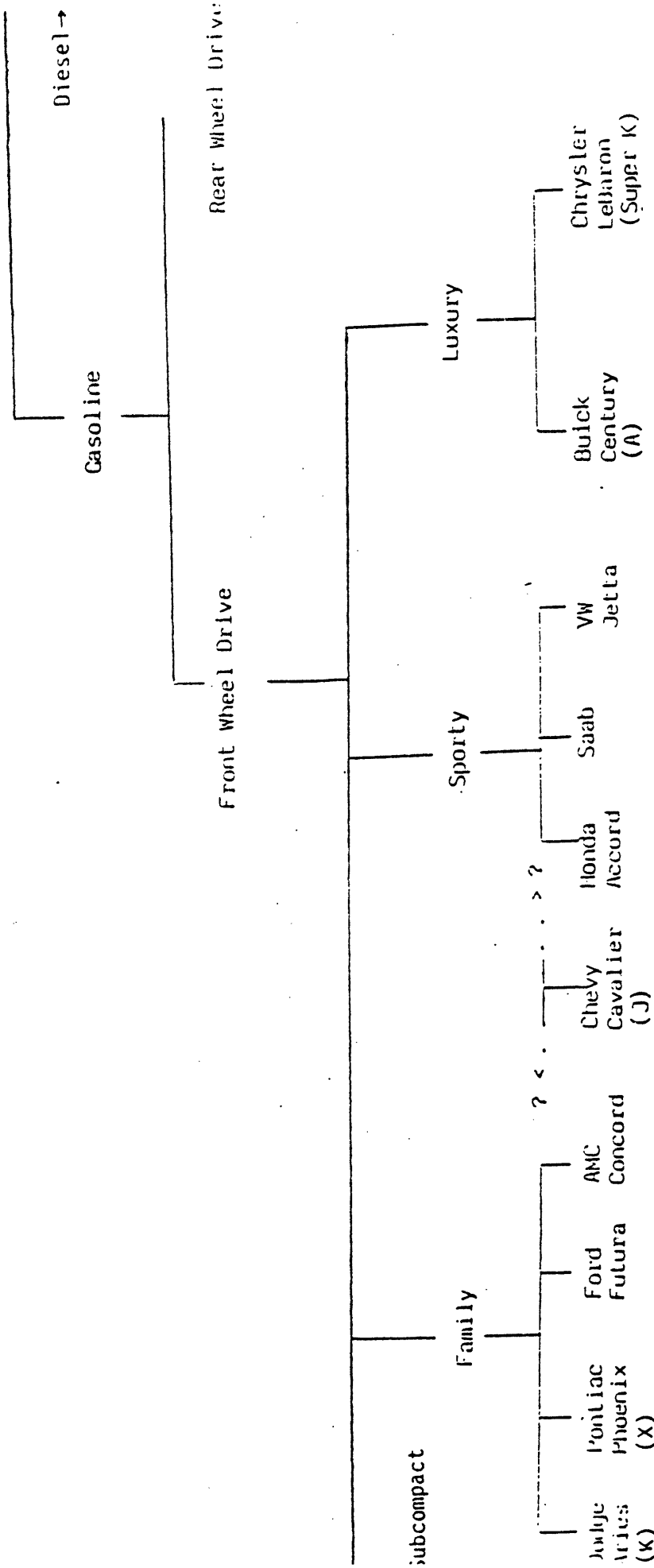


Figure 1: Hypothetical Structure of the Automobile Market. (Diesel, Rear Wheel Drive, and Subcompact branches may also have sub-branches. Letters in parentheses indicate that other, similar, brands are produced by the same manufacturer.)

Figure 1 is based on Consumer Reports (January/February 1982), but consider for a moment the managerial definition implicit in the above statement. We tend to say that the J-body cars are in the "sporty FWD" 'submarket' if they draw share from the Honda Accord, Saab, and Volkswagon Jetta; they are in the "family FWD" 'submarket' if they draw share from the Dodge Aries (K), Pontiac Phoenix (X), Ford Futura, and AMC Concord.

However, suppose the J-body cars were popular and capture, say a market share of 8.8% of the total number of FWD cars sold.<sup>4</sup> Suppose further that for this example, "family FWD" cars account for the majority of the sales of FWD cars, say 96%. Suppose the J-body cars captured all of the consumers who previously would have purchased sporty FWD imported cars (4% of the total market) and 5% of the consumers who previously would have purchased "family FWD" market (4.8% of the total market --  $5\% \times 96\%$ ). Do we classify the J-car as a "family FWD" car because it draws more consumers from that submarket than the "sporty FWD" submarket? Do we classify the J-car as a "sporty FWD" car because it draws all of the "sporty FWD" sales and only a small fraction of the "family FWD" sales? Do we classify it as neither? Or, both? The choice depends on managerial need and how we represent that need through our definition of market structure.

In this paper we choose our definition to reflect what we believe is the key message portrayed by a grouping of products into submarkets. Suppose a macro-market, e.g. "FWD automobiles" has been defined and suppose we want to ask whether that macro-market contains any submarkets, e.g., "sporty FWD" and "family FWD".

We must first contemplate an ideal unstructured market. In such an unstructured market, the manager would not expect his product to draw customers equally from all existing products. He would expect popular products to be hurt more than unpopular products, that is, he would expect to draw more customers from high share (of the macro-market) products than from low share products. Thus, we define an ideal

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<sup>4</sup> This example is purely illustrative. It does not reflect actual automobile shares.



unstructured market as a market in which a new product draws its share from existing products in proportion to the market shares of the existing products. In such an unstructured market, the manager would have no need to group products into submarkets since he could predict competitive impact by simply knowing the market shares of all existing products.

On the other hand, if his product were such that it hurt specific identifiable products more so than market shares would predict, he would want to predict this when formulating his market strategy. In our example, the Chevy Cavalier draws more customers from sporty FWD cars than would be predicted by a market share of 8.8%. A market structure grouping should convey this message to the manager. In other words, a market structure (as defined here) tells the manager that he will affect products grouped with his product more so than would be predicted by an ideal unstructured market.

This definition, in our opinion, captures the key managerial requirement of a market partition, that is, that the partitioning of the market into submarkets explains more about consumer behavior than is apparent to the manager from the unpartitioned market. According to this definition, the J-body cars would be classified as "sporty FWD" because they draw all of the previous "sporty FWD" share whereas their market share would predict that they would draw only 8.8% of the "sporty FWD" market. The above definition would reject classification of the J-body cars as "family FWD".

This definition is useful for conceptualization of the managerial issue. But we also want to identify competitive structure before a new product is introduced. Thus we will also operationalize the definition with respect to product deletion. In particular,

"A market is defined by a series of submarkets, if, when a product is deleted from a submarket, its former consumers are more likely to buy again in that submarket than would be predicted by market shares."

A moment's reflection reveals that both definitions capture essentially the same phenomena. We find the second definition easier to operationalize. Both relate directly to the impact of managerial changes in the composition of the product line and to the definition of competition.

Finally, we note that other marketing scientists may define structure in other ways depending upon the managerial problems they face and depending upon what they wish to portray to the manager. We have chosen a specific definition which we feel has intuitive appeal. It is an aggregate definition because the managerial issue it addresses is related to aggregate strategy.

What follows is a deductive mathematical analysis based on this definition.

#### STATISTICS

Our approach to statistical testing for the presence of submarkets is based on classical hypothesis testing. We formulate a null hypothesis that reflects the existence of no competitive submarkets. Next we hypothesize a structure. Each hypothesized structure predicts how the market will behave when a product is deleted. To be retained (escape rejection), the hypothesized structure must explain product deletion probabilities better (statistically) than the null hypothesis of no structure. As is the case with all statistical hypothesis testing (e.g. Green and Tull 1978; Morrison 1976), we may:

- (1) retain only one structure as better than the null hypothesis,
- (2) fail to reject the null hypothesis relative to each structure, or
- (3) retain more than one structure.

Managerial action in cases (1) and (2) is clear. We address later, what to do in case (3).

We begin with a model to analyze market behavior and formalize what we mean by "no structure."

#### Aggregate Constant Ratio Model

In order to use the above definition, we must state mathematically what we mean by "would be predicted by market shares". To give this definition rigor we use the aggregate constant ratio model (ACRM). See discussions by Tversky and Sattath (1979, p. 552) and Bell, Keeney, and Little (1975).

Basically ACRM is an aggregate version of what is known as Luce's axiom (1959). However, we caution the reader that ACRM is not an aggregation of individuals who themselves obey an individual level constant ratio model (CRM). ACRM is purely a statement about how an aggregation of individuals behave. It is relevant for our analysis because, like our definition, it applies to the aggregate behavior of a group of consumers. For further discussion and examples see Tversky and Sattath (1979, pp. 552-554).

According to ACRM there exist scale values,  $m_j$ , for every product,  $j$ , in the market. Then, for any submarket (set of products),  $\underline{A}$ , the market share,  $P(j|\underline{A})$ , of product  $j$  in that submarket is given by:<sup>5</sup>

$$P(j|\underline{A}) = \frac{m_j}{\sum_{k \in \underline{A}} m_k} \quad (1)$$

where the denominator is simply the sum over all products in the submarket,

A. If  $\underline{A}$  is the total market,  $\underline{T}$ , we normalize such that  $\sum_{k \in \underline{T}} m_k = 1$ .

In this case,  $m_j$  also becomes the market share because  $P(j|\underline{T}) = m_j$ .

To implement our definitions we introduce some simplified notation. The reader will note that sets and indices of sets are denoted by underlined **boldface type**.

$P_i(j)$  = the overall market share of product  $j$  when product  $i$  is deleted,  $i \neq j$ .

$\underline{s}$  = a set of products, called submarket  $\underline{s}$ .

$P_i(\underline{s})$  = the market share of submarket  $\underline{s}$  out of the total market when product  $i$  is deleted,

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<sup>5</sup>Equation 1 is similar in structure to that derived by Jeuland, Bass, and Wright (1980, p. 262) under the assumptions that (a) individuals are described by a multinomial process, (b) the parameters of the process are distributed via a Dirichlet distribution across individuals, and (c) brand choice and purchase timing are independent. Thus, our statistical tests can be utilized to examine the market boundaries used with the multinomial-Dirichlet brand switching model. If 'no structure' is rejected at the market level but applies (is not rejected) at the submarket level, then the multinomial-Dirichlet model may be appropriate for each submarket but not for the aggregate market.

From equation (1) and the above definitions we have:

$$P_i(j) = m_j / \sum_{k \neq i} m_k = m_j / (1 - m_i) \quad (2)$$

$$P_i(\underline{s}) = \sum_{\substack{j \in \underline{s} \\ j \neq i}} P_i(j) = \left( \sum_{\substack{j \in \underline{s} \\ j \neq i}} m_j \right) / (1 - m_i) \quad (3)$$

Where the sum in equation 3 is over all products in submarket  $\underline{s}$  except when  $i$  is in submarket  $\underline{s}$ , in which case we delete  $i$  from the sum.

For example, if the scale values ( $m_j$ ) of products 1, 2, 3, and 4 were each .25, then  $P_1(2) = P_1(3) = P_1(4) = .25 / (1 - .25) = .333$ . If the first submarket ( $\underline{s} = 1$ ) were products 1 and 2, and the second submarket ( $\underline{s} = 2$ ) were products 3 and 4, then  $P_1(\underline{s} = 1) = P_1(2) = .33$  and  $P_1(\underline{s} = 2) = P_1(3) + P_1(4) = .67$ .

Equations (2) and (3) predict what the new shares will be under the null hypothesis for any specific groupings of products into submarkets. We now compare these aggregate predictions of ACRM to the observed behavior of consumers. Let:

$n_i$  = number of consumers who choose product  $i$  when all products are available.

$n_i(j)$  = the number of consumers out of  $n_i$ , who formerly chose product  $i$  but who now choose product  $j$  when product  $i$  is deleted from the market.

$n_i(\underline{s})$  = the number of consumers, out of  $n_i$ , who formerly chose product  $i$ , but who now choose a product from submarket  $\underline{s}$ . (Product  $i$  is no longer available.)

If there were no market structure, then ACRM would apply for any grouping of products and we would expect the observed frequencies of purchases to satisfy:<sup>6</sup>

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<sup>6</sup>Note that if we consider the entire market, that is, if  $\underline{s} \equiv \underline{T}$ , then equation (5) is an identity since  $n_i(\underline{T}) \equiv n_i$  and  $P_i(\underline{T}) \equiv 1.0$ . However, for any  $\underline{s} \neq \underline{T}$ , equation (5) becomes an empirically testable statement.

$$n_1(j)/n_1 \approx P_1(j) \quad (4)$$

$$\text{and } n_1(\underline{s})/n_1 \approx P_1(\underline{s}) \quad (5)$$

where the approximate equality ( $\approx$ ) is due to sampling error for finite  $n_1$ .

Now let  $\hat{P}_1(\underline{s}) = n_1(\underline{s})/n_1$  be the estimated probability of buying in set  $\underline{s}$  when product 1 is not available (based on those consumers who previously chose product 1). Then, according to our definition of competitive structure, a structure defined by a specific set of submarkets exists when switching is greater in each submarket than would be predicted by "no structure". Thus, for each submarket,  $\underline{s}$ , in the structure we expect:

$$\hat{P}_1(\underline{s}) \geq P_1(\underline{s}) \quad \text{if } i \text{ is in } \underline{s} \quad (6)$$

$$\hat{P}_1(\underline{s}) \leq P_1(\underline{s}) \quad \text{if } i \text{ is not in } \underline{s} \quad (7)$$

where approximation is again due to sampling errors. Equations (6) and (7) are now a mathematical interpretation of our definition of competitive structure.

Consider the above example, where  $n_1 = 100$ ; then if  $n_1(\underline{s} = 1) = 34$  ( $\hat{P}_1(\underline{s} = 1) = .34$  vs.  $P_1(\underline{s} = 1) = .33$ ) and  $n_1(\underline{s} = 2) = 66$  ( $\hat{P}_1(\underline{s} = 2) = .66$  vs.  $P_2(\underline{s} = 2) = .67$ ) we would probably not reject the hypothesis of no structure. On the other hand, if  $n_1(\underline{s} = 1) = 98$  ( $\hat{P}_1(\underline{s} = 1) = .98$  vs.  $P_1(\underline{s} = 1) = .33$ ) and  $n_1(\underline{s} = 2) = 2$  ( $\hat{P}_1(\underline{s} = 2) = .02$  vs.  $P_2(\underline{s} = 2) = .67$ ) we would probably reject ACRM as a model of the market and, according to equation (6) and (7), say that the hypothesized structure explains observed behavior in a way consistent with our definition.

Note that the criteria we use to test a submarket, that is, the inequalities in equations (6) and (7), depend upon how we define the submarkets. Thus, for "good" groupings into submarkets, equations (6) and (7) will hold, while for other potential groupings they may not hold and, hence, we would reject such groupings as not being superior to 'no groupings.'

If  $n_1$  were infinitely large then we could use equations (6) and (7) directly, however, for finite  $n_1$  we must recognize sampling errors.

## Normal Test

Our null hypothesis of no structure is that the population satisfies ACRM for any grouping, thus, if a consumer is drawn randomly from those who previously purchased product  $i$ , his probability of choosing from submarket  $\underline{s}$  is  $P_i(\underline{s})$ . If successive draws are independent, then this creates a binomial process for  $n_i(\underline{s})$  with mean,  $n_i P_i(\underline{s})$ , and variance,  $n_i P_i(\underline{s})(1-P_i(\underline{s}))$ . If  $n_i$  is sufficiently large, then the Central Limit Theorem applies and the distribution of  $n_i(\underline{s})$  is given by:<sup>7</sup>

$$n_i(\underline{s}) \sim N[n_i P_i(\underline{s}), n_i P_i(\underline{s})(1-P_i(\underline{s}))] \quad (8)$$

where  $\sim N[\mu, \sigma^2]$  means distributed as normal with mean,  $\mu$ , and variance,  $\sigma^2$ .

We define "success" with respect to the set  $\underline{s}$  rather than with respect to the product  $j$  since that is consistent with our qualitative definition. If the  $m_j$  are known for all products in the total market, then the  $P_i(\underline{s})$  are given by equation 3 and the statistical tests are standard one-tail Z-tests based on inequalities (6) and (7).

Note that our null hypothesis does not assume homogeneous CRM consumers, but is rather an aggregate statement. Heterogeneity is a subtle, complex issue. As illustrated later in this paper (See Tables 1 and 2), it is possible to have every consumer satisfy CRM yet have market structure. For example, consider a market where 50% of the consumers satisfy CRM but with probabilities favoring sporty FWD cars and 50% satisfy CRM but with probabilities favoring family FWD cars, such a market clearly has structure.

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<sup>7</sup>See Drake (1967, pp. 212-221). A good rule of thumb is  $n_i < 20$ . Note also that the DeMoivre-Laplace Central Limit Theorem requires (suppress  $\underline{s}$ ) that  $n_i P_i > 3\sigma_i$  and  $n_i(1-P_i) > 3\sigma_i$  where  $\sigma_i = [n_i P_i(1-P_i)]^{1/2}$ . If  $n_i$  is not sufficiently large to assure that the inequalities are satisfied, we can replace equation (8) with a Poisson approximation, as discussed in Drake (1967, p.220).

However, it is easy to show that equation 8 is an upper bound for heterogeneous CRM consumer and exact for homogeneous CRM consumers<sup>8</sup>, thus we feel equation 8 is an appropriate variance formula to represent our aggregate definition. That is, it is appropriately conservative and tells us when  $n_i$  is large enough to have statistical confidence in the inequality comparisons of equations (6) and (7). For more indepth discussion of ACRM versus heterogeneous CRM consumers, see Tversky and Sattath (1979).

### Summary Statistics

A one-tail Z-test based on equation (8) compares for each product,  $i$ , a hypothesized market structure to the null hypothesis of no structure. Because our primary concern is the submarket,  $\underline{s}$ , that contains product  $i$ , it is useful to define the following summary statistics:

$$n(\underline{s}) \equiv \sum_{i \in \underline{s}} n_i(\underline{s}) \quad (9)$$

$$n^* \equiv \sum_{\underline{s} \in \underline{T}} n(\underline{s}) \quad (10)$$

where the first sum is over all products in  $\underline{s}$  and the second sum is over all submarkets in the total market,  $\underline{T}$ .

It is easy to write down the distribution of  $n(\underline{s})$ . Remember  $n_i$  is a different sample than  $n_j$  for  $i \neq j$ ; thus the terms in equations (9) and (10) are independent normal random variables. Therefore, the means, or the variances, are sums

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<sup>8</sup>Suppressing the argument  $\underline{s}$  and the subscript  $i$  and indexing consumers by  $c$ ,

$$\begin{aligned} \text{var}(n) &= \sum_c P_c (1 - P_c) = n\bar{P} - \sum_c P_c^2 = n\bar{P} - \left[ \sum_c (P_c - \bar{P})^2 + n\bar{P} \right] \\ &= n\bar{P} - n\bar{P}^2 - \sum_c (P_c - \bar{P})^2 \leq n\bar{P} - n\bar{P}^2 = n\bar{P}(1 - \bar{P}). \end{aligned}$$

of the means, or the variances, in equation 8.<sup>9</sup>

These summary statistics can also be re-expressed as proportions

$$\hat{P}(\underline{s}) = n(\underline{s}) / \sum_{i \in \underline{S}} n_i \quad (11)$$

and

$$\hat{P}^* = n^* / \sum_{i \in \underline{T}} n_i \quad (12)$$

The corresponding weighted ACRM proportions are

$$P(\underline{s}) = \sum_{i \in \underline{S}} P_i(\underline{s}) n_i / \sum_{i \in \underline{S}} n_i \quad (13)$$

and

$$P^* = \sum_{\underline{s} \in \underline{T}} (P(\underline{s}) \sum_{i \in \underline{S}} n_i) / \sum_{i \in \underline{T}} n_i \quad (14)$$

These aggregate statistics are useful in communicating the results of an evaluation of the overall competitive structure. (Actual statistical tests are made with respect to  $n(\underline{s})$  and  $n^*$ .) The empirical case represented at the end of this paper demonstrates the use of these summary measures of competition.

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<sup>9</sup>In particular,  $n(\underline{s}) \sim N[\sum_{i \in \underline{S}} n_i P_i(\underline{s}), \sum_{i \in \underline{S}} n_i P_i(\underline{s})(1-P_i(\underline{s}))]$  and

$n^* \sim N[\sum_{\underline{s} \in \underline{T}} \sum_{i \in \underline{S}} n_i P_i(\underline{s}), \sum_{\underline{s} \in \underline{T}} \sum_{i \in \underline{S}} n_i P_i(\underline{s})(1-P_i(\underline{s}))].$



## TESTING PROCEDURE

The previous section provides a means to turn our managerial definition into a statistical statement. Figure 2 summarizes one way in which a marketing scientist can use these statistics to help a manager identify a competitive market structure. We illustrate our procedures later when we apply our methods empirically.

The first step is to generate hypotheses about the nature of submarkets. Exploratory analysis is useful. Applying perceptual mapping or clustering procedures to usage data (e.g. Srivastava, Shocker, and Day, 1978; Day, Shocker and Srivastava, 1979; and Srivastava, Leone and Shocker, 1981) could yield hypotheses on submarkets characterized by use occasions. Applying analytical procedures to switching data (e.g. Kalwani and Morrison, 1977, or Rubinson, Vanhonacker and Bass, 1980) could generate hypotheses characterized by product attributes (e.g. brand or form). Clustering of individual attributes could define user characteristics or hypotheses on competitive sectors (e.g. Frank, Massy and Wind, 1972). Managerial judgement may also be utilized to formulate competitive hypotheses.

Next specific data is obtained to test each hypothesized grouping of products into submarkets. Experimentally controlled forced switching data is one way to estimate directly the forced switching frequencies,  $n_1(j)$ . We describe others.

Next the hypothesized groupings ( $H_A$ ) are tested statistically. Since each hypothesized grouping will imply different sets of inequalities as per equations (6) and (7), each hypothesized grouping is tested independently. (The inequalities vary because  $g$  varies.) In some cases, one can nest hypotheses, i.e., define ACRM on a submarket and test sub-submarkets. If a hypothesis is not significantly better than no structure at, say the 10% level, it is eliminated. After the hypotheses are tested, all or some may have been eliminated.

Before selecting the competitive structure for managerial purposes, we may wish to explore for the existence of "compound structures" in the definition of the

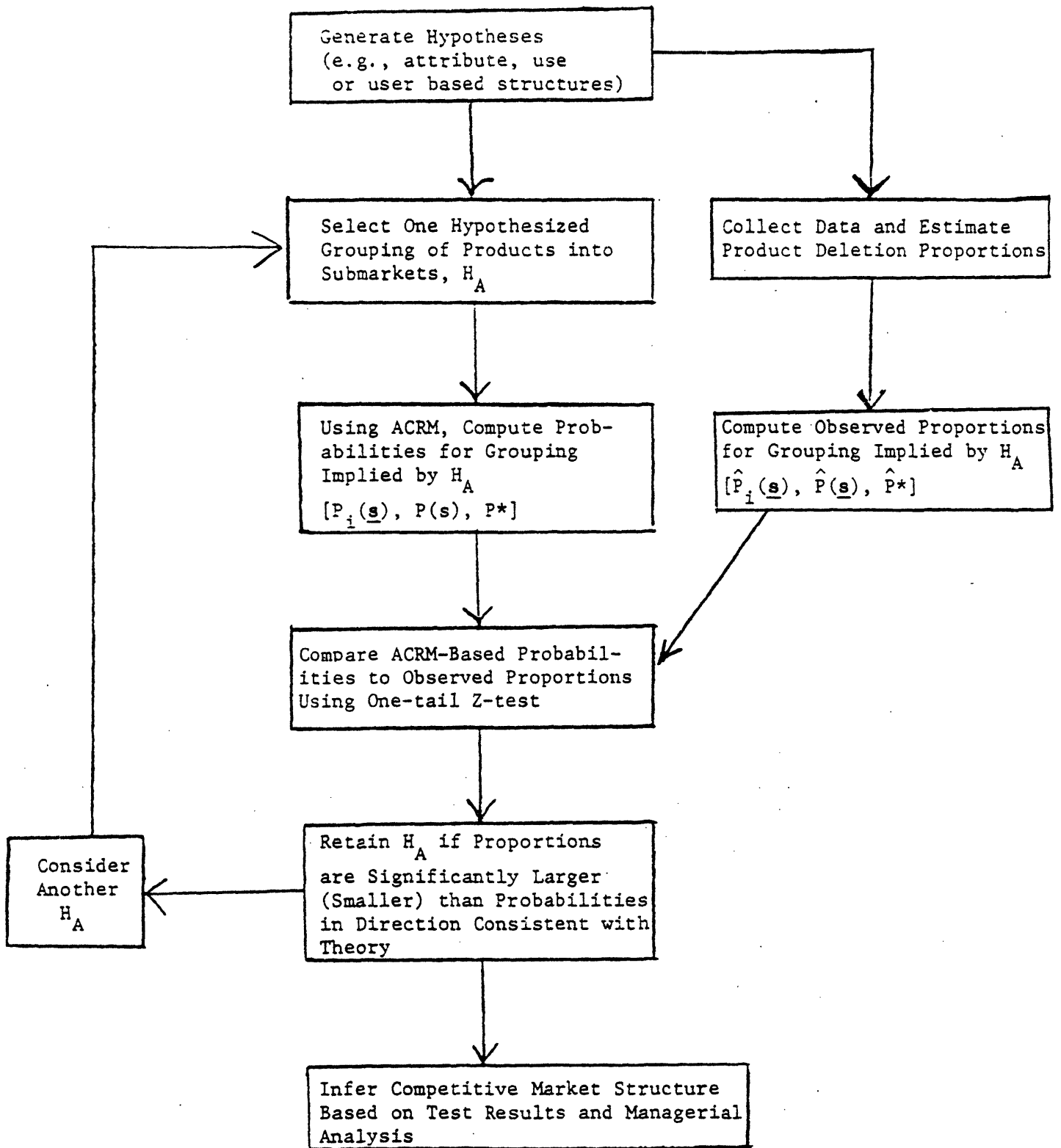


FIGURE 2: TESTING PROCEDURE

submarkets.<sup>10</sup> That is, we explore whether consumers can be grouped such that each group is characterized by a different competitive definition (e.g. people over 50 years of age may see the coffee submarkets as decaffeinated and caffeinated, while those under 50 may see the submarkets as ground and instant).

Finally we select the best structure for managerial analysis. If no test rejects the hypothesis of no structure in favor of a specific hypothesized structure, we use the unstructured description. If only one of the hypothesized structures is significant, it is subjected to further evaluation. Because alternative hypotheses have been tested, care should be taken in interpreting the significance level so as to not exploit random error. We recommend re-testing of the chosen hypothesis with saved data or through convergence from separate measures of switching. If the convergent analysis or retest is consistent, the hypothesis may be adopted for managerial analysis. It is clear at this point that the generation of hypotheses must span the relevant set of possibilities if the hypothesis testing procedure is to identify the best competitive definition.

A more difficult case exists if more than one of the attribute, use, or user hypotheses is significant. In this case we rely on managerial judgement. The significance levels of the alternative hypotheses and the results of convergent analysis are useful inputs to the decision, but the final choice will also reflect managerial experience.<sup>11</sup> Additional forced switching studies also may be undertaken to collect

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<sup>10</sup>We have chosen the words "compound structure" to avoid confusion with heterogeneity which is usually taken to mean a continuous probability distribution of varying CRM probabilities. We seek to identify different groups of consumers such that the null hypothesis of no structure is rejected within each consumer group, but for which the hypothesized structure varies across consumer groups.

<sup>11</sup>We avoid the temptation to choose the highest Z-score since that would be analogous to choosing a regression based on the highest F-score. Such procedures exploit random error. The Z-score provides a guide to the decision but should not be the only criterion. In making a choice it is useful to recognize that the variance of the summary statistic,  $n^*$ , is summed over all products and hence depends upon the total sample size. The variances of the more specific submarket statistics,  $n(\underline{s})$  are summed only over those products in  $\underline{s}$  and hence depend upon  $\sum_{\underline{s}} n_1$ . Since the power of a Z-test increases with sample size, the likelihood of falsely retaining the null will be similar across alternative hypotheses when using the summary statistics. On the other hand, the likelihood of false retention will be larger for "smaller" submarkets, i.e., smaller  $\sum_{\underline{s}} n_1$ . Thus, care should be exercised when interpreting Z-statistics for small submarkets or when interpreting the Z-statistics testing  $n_1(\underline{s})$  for small share brands.

more evidence for selection of the best competitive description. This would be particularly true if a compound structure were obtained from the exploratory analysis.

Given a competitive structure, managerial analysis could be conducted to assess opportunities for new products, the advisability of dropping products to consolidate product offerings, or to answer questions such as "who is my competition?", or "what universe of products do we use to calculate market share for strategic purposes?"

In the following sections, we describe and illustrate the steps in the testing process and present an empirical application. We begin with data collection and estimation.

#### DATA COLLECTION AND ESTIMATION PROCEDURES

The construct upon which the Z-statistics are based is  $n_i(j)$ , the number of consumers who buy product  $j$  when product  $i$  is deleted from the market. In this section we discuss how one might measure or estimate  $n_i(j)$ . Greater detail on the actual measures are given in the empirical case.

##### Forced Switching

The most natural method of observing  $n_i(j)$  is to first observe the product that an individual most prefers then place him in a choice situation in which his preferred product has been removed from the choice set. We call this experiment forced switching. In our empirical case we use a simulated store to observe each consumer's choice from the appropriately modified choice sets.

##### Preference Rank

An alternative data collection procedure is to ask consumers to rank order the products they would consider in terms of preference. For a given individual we identify product  $i$  as their first ranked product and product  $j$  as their second ranked product.

Then,  $n_i(j)$  is estimated as the number of consumers who rank product  $i$  first and product  $j$  second. To the extent that rank order preference reproduces actual choice, this measurement will provide an estimate of the true  $n_i(j)$ . Another variant of this approach is to calculate the fraction of people who last purchased their most preferred brand, 2nd most preferred brand, etc. This vector can then be applied to each consumer's rank order brand preferences to provide a probability of purchase for each product in their consideration set. These can be aggregated across consumers to estimate  $n_i(j)$ .

### Logit/Preference Intensity

Suppose that we can estimate for each consumer,  $c$ , his probabilities,  $P_c^-(i)$ , of choosing product  $i$ . Suppose we can estimate  $P_c^-(i)$  for all products in consumer  $c$ 's consideration set. Suppose further that all individual consumers satisfy CRM.

Then, for consumer,  $c$ , the probability,  $P_{ci}^-(j)$ , that he chooses product  $j$  from the set of products in which product  $i$  has been deleted is given by:

$$P_{ci}^-(j) = P_c^-(j) / [1 - P_c^-(i)]. \quad (15)$$

where  $P_c^-(j) = 0$  if  $j$  is not in  $c$ 's consideration set.  $n_i(j)$  is then obtained by simple aggregation of conditional probabilities,<sup>12</sup> i.e.:

$$n_i(j) = \sum_c P_{ci}^-(j) P_c^-(i) \quad (16)$$

$$n_i = \sum_c P_c^-(i) \quad (17)$$

The first term is the probability that  $c$  chooses  $j$  second given that he chose  $i$  first; the second term is the probability that  $c$  chooses  $i$  first.

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<sup>12</sup> $n_i(j)$  based on equation (16) is an estimate of the  $n_i(j)$  that would be observed empirically. Although the sampling variance is different if  $P_c^-(i)$  varies across  $c$ , equation (8) still serves as an upper bound as shown in footnote 8. An alternative hybrid technique uses first preference to compute  $n_i$ . Then,  $n_i(j)$  is computed via equation (15) summed across all consumers who prefer  $i$ .

For example, suppose there are four products on the market and two consumers with the CRM probabilities shown in Table 1.

TABLE 1: EXAMPLE OF INDIVIDUAL PROBABILITIES

	<u>Consumer 1</u>	<u>Consumer 2</u>
Product 1	.4	.1
Product 2	.4	.1
Product 3	.1	.4
Product 4	.1	.4

For consumer 1,  $P_{11}(2) = .67$  ( $.67 = .4/(1-.4)$ ),  $P_{11}(3) = .17$ ,  $P_{11}(4) = .17$  and for consumer 2,  $P_{21}(2) = .11$ ,  $P_{21}(3) = .44$ ,  $P_{21}(4) = .44$ . The estimate of  $n_1(2)$  is then given by:

$$n_1(2) = P_{11}(2)P_1(1) + P_{21}(2)P_2(1) = (.67)(.4) + (.11)(.1) = .279$$

$$n_1 = P_1(1) + P_2(1) = (.4) + (.1) = .5$$

Thus,  $\hat{P}_1(2) = n_1(2)/n_1 = .558$ . The reader may wish to verify that we obtain an (estimated) forced switching matrix in Table 2 which suggests a definite market structure grouping products 1 and 2 as one submarket and products 3 and 4 as the other submarket. This structure occurs because of heterogeneity of preference within the individual CRM models. This example illustrates a case where, although each individual satisfies CRM, the market is structured.

TABLE 2: ESTIMATED FORCED SWITCHING MATRIX [ $P_i(j)$ ]

		<u>Product</u>			
		Product 1	Product 2	Product 3	Product 4
<u>Product:</u>	Product 1	-	.56	.22	.22
	Product 2	.56	-	.22	.22
	Product 3	.22	.22	-	.56
	Product 4	.22	.22	.56	-

There are many ways to estimate the individual CRM probabilities. The most common method, is the multinomial logit model with constant sum paired comparison preference data as the explanatory variables. (Actual choice is the dependent quantal variable in the estimation.) This method assumes that CRM is true over each individual's consideration set. For more details see McFadden (1980), Silk and Urban (1978), and Urban and Hauser (1980, Chapter 11).

## Consideration Sets

In some cases the only data available may be an indication of which product is preferred and which products each consumer would consider purchasing. Sometimes we only know the consideration set information. Intuitively we expect that these consideration sets carry much information about what products group together in submarkets. Suppose that consumer  $c$  evokes  $e_c$  acceptable products where  $e_c \geq 2$ . Then if we assume consumers are equally likely to buy any product within their consideration sets, we get:

$$P_c(i) = \begin{cases} 1/e_c & \text{if } i \text{ is considered by consumer } c \\ 0 & \text{if } i \text{ is not considered by consumer } c \end{cases} \quad (18)$$

We can use equation (18) with equations (15), (16), and (17).

Alternatively, if we can identify the first preference product, we can use a hybrid approach. That is, we compute  $P_{c_i}(j)$  with equations (15) and (18), but estimate  $n_i(j)$  by summing over only those consumers who prefer product  $i$ .  $n_i$  becomes the number of consumers who prefer product  $i$ .

## Convergence

The forced choice procedure measures the aggregate proportion of respondents switching to product  $j$  when  $i$  is not available, while the preference rank, Logit, and consideration set procedures calculate these probabilities based on individual choice probabilities. A convergent approach can be used when both experimentally forced choices and individual probabilities are available. The probability of buying again in the submarket ( $P_i(\underline{s})$ ) can be estimated by each method and the statistical adequacy assessed. If both methods agree, confidence is increased. If they disagree, an examination of sources of bias may reveal, in a particular application, that one method is preferred. If no bias is identified, the probabilities ( $P_i(\underline{s})$ ) can be pooled across methods to obtain a combined assessment of the competitive structure (see application for more details).

## TESTING HYPOTHESES

In this section, we formulate and illustrate the statistical testing for hypotheses based on attributes, on use, and on user.

### Attribute Hypotheses

We begin with a simplified example of how the normal test would be applied to competitive structures based on attributes of automobiles. Suppose there are only seven specific models (or brands) of automobiles on the market; three diesels (Peugeot 505, Olds Cutlass, and VW Jetta) and four gasoline powered cars (Chevrolet Cavalier, Peugeot 505, Olds Cutlass, and VW Jetta). Table 3 represents a hypothetical data matrix for 100 consumers. The numbers in the first column represent the number of consumers who selected that model of automobile as their first choice,  $n_i$ . The numbers in the matrix represent the number of consumers whose first choice is the automobile designated by the row label and whose second choice is the automobile designated by column label,  $n_i(j)$ . For simplicity in this example, we assume the share of product  $i$  is  $m_i = n_i / \sum_1 n_i$ .

TABLE 3: HYPOTHETICAL FORCED SWITCHING MATRIX [ $\hat{P}_i(j)$ ]

		Diesel			Gasoline			
	$n_i$	Peugeot	Cutlass	Jetta	Cavalier	Peugeot	Cutlass	Jetta
<b>Diesel</b>								
Peugeot	10	-*	6	2	1	1	0	0
Cutlass	20	10	-	4	0	2	4	0
Jetta	15	7	2	-	1	1	1	3
<b>Gasoline</b>								
Cavalier	20	0	2	4	-	0	4	10
Peugeot	15	3	1	1	0	-	5	5
Cutlass	10	0	1	0	3	3	-	3
Jetta	10	0	0	2	1	5	2	-

\* No repeat purchased is allowed under the forced switching conditions.



Two alternative hypotheses for a competitive structure are "Diesel vs. Gasoline" as shown in Table 4a and "model-specific" as shown in Table 4b. Inspection of the data in Table 1 suggests that "diesel" consumers would stay with diesels and "gasoline" consumers would stay with gasoline powered automobiles if they were forced to switch from their most preferred car. But does this hypothesis hold up statistically?

TABLE 4: TWO ALTERNATIVE COMPETITIVE STRUCTURES

a) "Diesel" vs. "Gasoline"

<u>Diesel</u>		<u>Gasoline</u>
Peugeot		Cavalier
Cutlass		Peugeot
Jetta		Cutlass
		Jetta

b) "Model-specific"

<u>Peugeot</u>	<u>Cutlass</u>	<u>Jetta</u>	<u>Cavalier</u>
Gasoline	Gasoline	Gasoline	Gasoline
Diesel	Diesel	Diesel	

Table 5 uses the Z-statistics to test the alternative hypotheses. We use the notation  $\hat{P}_i(\underline{s}) = n_i(\underline{s})/n_i$  to make the comparison to  $P_i(\underline{s})$  where  $\underline{s}$  is the submarket that contains product i. Since the cutoff level for a one-tailed significance level of .10 is 1.28, Table 5 suggests that the "Diesel vs. Gasoline" market structure is significantly better than no structure at the .10 level for all seven products. On the other hand, the Z-tests clearly suggest that the model-specific hierarchy is not significantly better than a hypothesis of no structure. The aggregate tests of  $\hat{P}(\underline{s})$  versus  $P(\underline{s})$  also indicate the "Diesel vs. Gasoline" partitioning of the market is significantly better than no structure.

In the above illustration we have used the statistics to test two alternative hypotheses about market structure. One was retained; the other rejected. However, in some applications more than one hypothesis could be retained. This would be analogous to comparing two regressions based on different variables. It is tempting to accept the structure with the highest Z-statistic (the regression with the highest F), but in doing

so one must realize it is not a rigorous use of the statistic. We recommend instead that one exercise the same cautions normally exercised when choosing the "best" regression model. See for example, discussions in Drake (1967), Green and Tull (1978), and Morrison (1976).

TABLE 5: TEST STATISTICS FOR ALTERNATIVE STRUCTURES

( $n_i$  is the number in parentheses)

	"Diesel vs. Gasoline"			"Model-specific"		
	$\hat{P}_i(s)$	$P_i(s)$	Z	$\hat{P}_i(s)$	$P_i(s)$	Z
Peugeot Diesel(10)	.80*	.39**	2.66***	.10	.17	-.58
Cutlass Diesel (20)	.70	.31	3.77	.20	.13	.93
Jetta Diesel (15)	.60	.35	2.02	.20	.12	.95
Cavalier Gasoline (20)	.70	.44	2.34	-	-	-
Peugeot Gasoline (15)	.67	.47	1.55	.20	.12	.95
Cutlass Gasoline (10)	.90	.50	2.53	.10	.22	-.92
Jetta Gasoline (10)	.80	.50	1.89	.20	.17	.25
	$\hat{P}(s)$	$P(s)$	Z	$\hat{P}(s)$	$P(s)$	Z
Aggregate Test	.72	.41	6.10	.18	.15	.84

\*Note .8 equals fraction of Peugeot choosers who select a diesel  
 Cutlass or Jetta (.8 = (2+6)/10 from Table 1)  
 \*\*by equation 3 (.39 = (.2+.15)/(1-.1))  
 \*\*\*(.8-.39)/[.39(1-.39)/10]<sup>1/2</sup>

One can use the Z-statistic in theory testing mode as illustrated above or in exploratory mode where one uses the statistic to search for the "best" managerial competitive structure. In the latter case we suggest that the chosen hypothesis then be tested with either convergent methods, holdout samples, or both.

### Use Hypotheses

Some firms position their products for particular uses and such use categories can become the basis for defining the structure of competition. For example, in the home cleaner market, cleaning the kitchen may be one submarket and cleaning the bathroom another. We would find a competitive structure defined by use if one set of products tend to be used for the kitchen and another set of products for the bathroom.

We can proceed to test use hypotheses in one or two ways. The first way is to collect separate data for each use and follow the procedure in Figure 2 to identify a structure within each use. Since this procedure is a simple extension, we need not illustrate it here.

The second procedure is more complex. Suppose the manager wishes to assign products to use groupings and then test those groupings to determine if consumers stay within groupings when choosing another product for the same use.

The first procedure is a set of multiple structures where, for each use, all products are assigned to submarkets. The second procedure is a single structure where each product is assigned to one group and each group corresponds to a use. We illustrate here the second procedure.

In the second procedure we begin by uniquely assigning products to submarkets. Let  $n_{iu}$  = the number of consumers who use product  $i$  for use  $u$ .

One reasonable assignment rule is to assign product  $i$  to the use submarket  $\underline{u}$  which maximizes  $n_{iu}$ .

We now consider deleting products. If we delete product  $i$  (which is assigned to use submarket,  $\underline{u}$ ) and ask consumers to consider only use  $u$ , then we would hope that consumers are more likely to choose again from products assigned to use submarket  $\underline{u}$  than would be predicted by market share.

Testing this hypothesis mathematically is very similar to testing a market structure defined by brands or by product characteristics. There is one important difference. Suppose there are four cleaning products on the market, Ajax, 409, Top Job, and Mr. Clean. Suppose the above rule assigns Ajax and 409 to bathroom cleaning and assigns Top Job and Mr. Clean to kitchen cleaning. To test a hypothesized market structure we must decide how to deal with consumers whose choice violates the assignment rule. For example consumers who choose Ajax for kitchen cleaning are misclassifications according to our assignments.

For such consumers we chose to remove not Ajax but their most preferred "kitchen" product, where "kitchen" is defined by the procedure described above. In this case, we remove their most preferred of Top Job or Mr. Clean. Note that for all consumers "kitchen" or "bathroom" products are defined by the market structure we are testing. Such "misclassified" consumers will of course still prefer Ajax when either Top Job or Mr. Clean are removed because they preferred Ajax when all products were available. Thus such consumers will be counted as evidence against the hypothesized structure. This will assure that any statistics we compute will be appropriately conservative. We now turn these verbal statements into mathematical statements. These definitions are similar to those defined earlier except that we now condition on usage,  $u$ . Let:

$n_{iu}(j)$  = the number of consumers who would purchase product  $i$  from the set of products designated for use  $u$ , and who would purchase product  $j$  for use  $u$  when product  $i$  is deleted. Note that  $n_{iu}(j)$  is only defined for products,  $i$ , contained in usage product set,  $\underline{u}$ .

$n_{iu}(\underline{u})$  = the number of consumers who would purchase product  $i$  from the product set  $\underline{u}$  for use  $u$ , and who would purchase a product, other than  $i$ , from usage product set  $\underline{u}$  if  $i$  were deleted. Note that  $n_{iu}(\underline{u}) = \sum_{j \in \underline{u}} n_{iu}(j)$ .

Then, if there is no structure and ACRM holds independent of use, we expect

$$\hat{P}_{iu}(\underline{u}) \approx P_{iu}(\underline{u}) \quad (19)$$

where  $\hat{P}_{iu}(\underline{u}) = n_{iu}(\underline{u})/n_{iu}$  and  $P_{iu}(\underline{u}) = (\sum_{\substack{j \in \underline{u} \\ j \neq i}} m_j)/(1-m_i)$

where the summation is over all products assigned to use submarket,  $\underline{u}$ , except product  $i$ . If submarkets exist we would expect

$$\hat{P}_{iu}(\underline{u}) > P_{iu}(\underline{u}) \text{ for } i \text{ in } \underline{u}$$

$$\hat{P}_{iu}(\underline{u}) < P_{iu}(\underline{u}) \text{ for } i \text{ not in } \underline{u}$$

The contribution of equation 19 is the definition of use structure relative to no structure. Statistical testing follows the same procedures as before. We can estimate  $n_{iu}(\underline{u})$  by any of the four methods discussed above, if we incorporate the following modifications:

- (i) data (choice, preference, logit probabilities, consideration) are collected for each use, and
- (ii) forced switching is defined with respect to the first choice for each use.

In the above discussion we assigned products to submarkets based on their maximum use. Other assignment rules are possible. The association of products to use submarkets can also be made by factor analysis or cluster analysis of a matrix of the

probability of purchase of each product (j) for each use (u). (See Day, Shocker, and Srivastava, 1979 and Srivastava, Leone, and Shocker, 1981). In such a case the submarket set u would represent a composite of uses. Managerial assignment could also be used (e.g., products appropriate for breakfast versus those for lunch or for dinner). Note, however, in this composite use procedure we require each product to be assigned only to one submarket. Our first procedure which utilizes submarkets for use allows products to be in multiple submarkets. Consider now an example.

Table 6 is an illustrative hypothetical example for the cleaning market. There are two uses, kitchen and bathroom, and four products, Ajax, 409, Top Job, and Mr. Clean. The hypothesis we wish to test is that Ajax and 409 are bathroom products while Top Job and Mr. Clean are kitchen products. Note that kitchen forced switching is only defined with respect to kitchen products and that bathroom forced switching is only defined with respect to bathroom products.

TABLE 6: ILLUSTRATIVE EXAMPLE OF DATA TO TEST USE COMPETITIVE STRUCTURE

		<u>Bathroom Cleaning, <math>n_{iu}(u)</math></u>			
	$n_{iu}^*$	Ajax	409	Top Job	Mr. Clean
Ajax	250	-	200	25	25
409	150	100	-	25	25
Top Job				-	
Mr. Clean					-

		<u>Kitchen Cleaning, <math>n_{iu}(u)</math></u>			
	$n_{iu}$	Ajax	409	Top Job	Mr. Clean
Ajax		-			
409			-		
Top Job	200	25	25	-	150
Mr. Clean	200	25	25	150	-

\* Note:  $n_{iu}$  = the number of consumers who would purchase product i for use, u.

If the market shares were .31 for Ajax, .19 for 409, .25 for Top Job, and .25 for Mr. Clean then we get the Table 7 values for testing the usage submarkets. In this example, a usage based competitive structure explains the data significantly better than a hypothesis of no structure.

TABLE 7: TEST STATISTICS FOR USAGE SUBMARKETS

	$\hat{P}_{iu}(\underline{u})$	$P_{iu}(\underline{u})$	Z
Ajax (250)	.80	.27	17.9
409 (150)	.67	.38	7.4
Top Job (200)	.75	.33	12.7
Mr. Clean (200)	.75	.33	12.7

	$\hat{P}^*$	$P^*$	Z
Aggregate Test	.75	.32	25.8

### User Hypotheses

Many researchers have hypothesized that a market is divided into submarkets by user characteristics. For example, one might hypothesize that heavy beer drinkers (more than 3 beers per day) tend to use Miller Lite, Coors, or Budweiser, while light beer drinkers (2 or fewer beers per day) tend to use Schlitz, Anheuser Busch Natural, or Miller High Life.

To test user based hypotheses about competitive structure we create test statistics analogous to either procedure described for usage based hypotheses. The only difference in the second procedure is that the rows of tables analogous to Table 6 are based only on the appropriate user group. For example, if a heavy user prefers Miller Lite, forced switching for him is with respect to Miller Lite. Forced switching with respect to Miller Lite is computed only with respect to heavy users. If a heavy user prefers Schlitz (a light user brand), forced switching is computed with respect to his first choice among heavy user brands. As in the usage statistics, these assignments assure that any statistics are appropriately conservative, e.g., the heavy user who prefers Schlitz will still prefer Schlitz thus providing evidence against a user based competitive structure.

Because the concepts of "use" and "user" are so similar we leave the details on "user" competitive structure to the reader. We note that to collect forced switching data for "use" or "user" competitive structures, it may be necessary to generate hypotheses based on judgment and/or prior research.

#### EXPLORATORY ANALYSIS OF COMPOUND STRUCTURE

All tests outlined above are tests of the aggregate competitive structure. Even the "user" competitive structure, which, at first glance, appears to be an assumption of compound structure, is a test of aggregate structure. Products are assigned to submarkets based on user attributes, but the market is described by one set of submarkets.

It is possible that different groupings of consumers could be characterized by different competitive structures. For example, perhaps heavy beer drinkers characterize competition by brand while light beer drinkers characterize competition by use occasions (e.g., with or without guests). If such a compound hypothesis is formulated then the statistical testing of that hypothesis is straightforward. Simply apply the statistical tests within each grouping of consumers.

Identifying groupings, that is, assigning consumers to groups, is more problematic. Our null hypothesis of no structure is at the aggregate level. So, if we attempt to apply our statistics at the individual level, we may structure the market inappropriately.

We propose two alternative heuristics and caution the reader that these heuristics are developed for exploratory use only and for input to the managerial selection of the best description. Any groupings identified with these statistics must still be subjected to aggregate testing within the grouping. To retain rigor, a different sample should be used for testing than the one used for exploration. The heuristics are given below. Each assumes that a prior set of possible competitive structures has been



identified. The raw data are based on the logit/preference intensity or evoked set equations for  $P_{ci}(j)$ , equations (15) and (18); each heuristic assigns consumers to one of the competitive structures. The rules can be applied for a target product,  $i$ , or for a weighted sum across  $i$ .

(1) Assign consumers to the competitive structure which maximizes  $P_{ci}(\underline{s})$  for the submarket which contains  $i$ .  $P_{ci}(\underline{s})$ , as well as the definition of  $(\underline{s})$ , will of course vary across alternative competitive structures.

(2) Compare individual level product deletion probabilities to market level product deletion probabilities computed via ACRM. Assign consumers to the competitive structure with the largest difference,  $P_{ci}(\underline{s}) - P_i(\underline{s})$ , for the submarket which contains  $i$ .

Applying these heuristics allows an exploratory diagnosis of the presence of different structures within subsets of the market.

#### MANAGERIAL ANALYSIS

The final step in our testing procedure is the managerial decision of selecting the best structure to describe the competitive relationship in the market. As we described above, the managerial selection is based on judgment aided by the statistical measures. Managerial analysis is best discussed by example. Thus, we illustrate these issues more fully by describing the application of our testing procedure to the market for coffee. We then discuss the managerial implications of six other applications to date.

## EMPIRICAL APPLICATION

### Hypothesis Generation

Past research on market structure indicates several bases for competition in the market for coffee: brand, product attribute, usage, and user characteristic.

The most obvious alternative is a brand structure (see Figure 3a). Discussion with managers and focus groups indicated a candidate product attribute description based on dividing the market into six groups:

1. ground caffeinated
2. ground decaffeinated
3. instant caffeinated regular
4. instant caffeinated freeze dried
5. instant decaffeinated regular
6. instant decaffeinated freeze dried

In formulating a competitive structure based on use occasion, we considered the results of focus group discussions on coffee consumption. From these groups, we identified over 40 possible use scenarios. These were grouped by judgment into nine classes that were felt to span the use environment at a meaningful level of detail.

They were:

1. to start the day/with breakfast,
2. between meals/daytime alone,
3. between meals with others,
4. with lunch,
5. with supper,
6. dinner with guests,
7. in the evening
8. to keep awake in the evening, and
9. on weekends

In a sample of 295 users of coffee, each respondent indicated the subset of the nine occasions that applied to them and identified the product they would evoke for each occasion (see data collection description below for more detail).

We aggregated the occasions into six classes to obtain reasonable sample sizes for each brand and occasion (see Appendix for the data). We factor analyzed the matrix of evoking of brands for specific uses (see Day, Shocker, and Srivastava 1979). Inspection of the matrix of consideration of brands for specific uses shows little variation in the proportions considering a given product across uses (see Appendix). The factor analysis resulted in one dimension of use where most uses loaded on that dimension ( $\Lambda_1 = 6.45$  for first factor,  $\Lambda_2 = .28$  for the second factor). Other analyses were conducted to identify possible use hypotheses. A diary panel record of uses of coffee was collected from the respondents. Each coffee serving was recorded for a one week period, along with a description of the occasion. 54 percent of the respondents used only one brand over all occasions in the week and 10 percent more used one brand for all uses until it ran out and then switched to a new brand for all subsequent uses. (See Laurent, 1978 for a more extensive discussion of this data.) The pantry check indicated 43 percent had only one container of coffee on hand and 60.6 percent had only one container open. There was no strong evidence of usage as the basis of product competition, but an alternative was formulated based on managerial priors and interpretation of the data above. This hypothesis was kept simple and was based on grouping the occasions into two classes by time of day of use (A.M. or P.M.). Brands were assigned to either the A.M. or P.M. group based on whether they were most heavily evoked for A.M. or P.M. use occasions (see Figure 3c).

A competitive structure alternative based on user characteristics was formulated by defining two submarkets based on user purchase rates (heavy--more than one purchase per two weeks, and light--one or fewer purchases per two weeks). Products were assigned to the submarkets where their consideration proportions were highest (see Figure 3d).

The four hypotheses indicated above were subjected to formal statistical testing, based on our product deletion criteria and statistics. We note that other managers and other researchers may have preferred different hypotheses and such hypotheses are viable candidates for future tests.

### Data Collection

In accordance with measurement procedures outlined above, 295 users of coffee (those who drink more than one cup of coffee/per day at home) were interviewed in Springfield, Massachusetts, and Indianapolis, Indiana in July and August of 1977. Respondents were interviewed after being recruited in a shopping mall and quotas were set to assure at least 50 respondents used each major type of coffee (ground/instant, caffeinated/decaffeinated, freeze dried). Evoked uses, the products considered for each use, and the last product used were identified for each respondent. Respondents indicated the brands they would consider using for each of the nine usage scenarios that applied to them. Preferences for products for each use were obtained on a seven point scale (extremely well-liked to very much disliked). Brands were rated on 12 product attribute scales. After providing demographic data and answering questions on coffee consumption, respondents were given an opportunity to purchase coffee for their most frequent use with a two dollar coupon they were given as compensation for participating in the interview. When the respondents reached the shelf, they found their first preference product "out of stock". Eighty-five percent of the people made a purchase in the lab and seventy percent noticed their favored brand was missing. At the close of

the lab phase, respondents were requested to participate in a usage panel in which they would record in a diary for a week each cup of coffee served at home (when, kind, brand, who present, how many cups, who prepared). Sixty percent returned a complete diary. Two weeks after the lab, respondents were called back to determine their home inventory of coffee (kinds, brand, open or not, size of package) and recent purchases. The panel and pantry check were conducted in this application to provide additional insight into the effects of the use situation on product choice. In most applications, the evoking and preference by use would be sufficient to determine if use was the best basis for structure in the market.

### Hypothesis Testing

We need to calculate the number of respondents who prefer product  $i$  and who will purchase in submarket  $s$  if brand  $i$  is deleted [ $n_i(s)$ ]. We do our initial testing based on the use of  $P_c^-(i)$  estimated by the preference rank data method outlined above.

The estimated values were good based on the information theoretic test<sup>13</sup> -- 80 percent of the total uncertainty was explained ( $U^2 = .795$ ). The standard deviation between actual and predicted market shares was .009. In this application, the McFadden, Train, Tye (1977) residual test indicated that the estimated probabilities [ $P_c^-(i)$ ] were not subject to error due to independence of irrelevant alternatives.<sup>14</sup> After comparisons based on preference rank probabilities, the laboratory store shopping data was used for validation.

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<sup>13</sup> $U^2$  is based on an information theoretic interpretation of the uncertainty explained by the individual level choice probabilities. The denominator is the uncertainty (entropy) that would be explained by "perfect information", the numerator is the uncertainty explained by the probabilistic model. Thus  $U^2 \leq 1$ . For derivations and examples see Hauser (1978).

<sup>14</sup>An addendum, available from the authors, describes the results of this test and an analysis with hierarchical logit and compound structure based logit.

The first alternative tested is a brand structure (see Figure 3a). The number of respondents,  $n_i$ , choosing brand  $i$  was based on first preference. The number of respondents buying again in that submarket when their most preferred brand was not available ( $n_i(j)$ ) was then calculated with equation 15 summed across consumer preferring brand  $i$ . Summary statistics were calculated by equations 9 through 12. The ACRM probabilities were calculated by equations 2 and 3 where market shares relevant to the sample were used as scale values ( $m_i$ ) and the aggregate ACRM proportions were calculated by equations 13 and 14. In general respondents were less likely to repeat purchase another variant of brand than the null hypothesis of no structure would predict (see Figure 3a). The brand structure does not adequately describe the market because the probabilities ( $\hat{P}$ ,  $\hat{P}(s)$ ) are low and the hypothesized brand structure is not significantly better than the null hypothesis of no structure. ( $Z$  less than zero for one tailed test).

A product attribute description of the market was developed by dividing the products into the six groups described above. This competitive structure was statistically superior to the null hypothesis of no structure in explaining observed market behavior (see Figure 3b). All of the submarket switching probabilities are significantly higher than the probabilities predicted by no structure at the five percent level. The overall test is significant at the one percent level ( $Z = 11.3$ ). Respondents were more likely to buy another brand of the product with a specific attribute than a model of no structure would suggest. For example .47 of those who have a first preference for a brand of caffeinated ground coffee would purchase another caffeinated ground coffee if their most preferred brand were not available. The "no structure" probability is .11 and this difference is significant at the one percent level ( $Z = 8.4$ ).

The third hypothesis is based on uses of coffees (see Figure 3c). The  $Z$ -statistics indicate that the use grouping is not significantly better than the null hypothesis of no structure. In fact, since the  $Z$ -statistics are negative they suggest qualitatively

that the use grouping in Figure 3c actually does worse than the hypothesis of no structure in explaining consumer behavior. Thus this submarket hypothesis does not appear to be a good overall basis for identifying the competitive structure of products in the market for coffee.

The user hypothesis tested here (see Figure 3d), likewise, does not appear to be a good overall basis for defining competitive structure in the market for coffee.

ALTERNATIVE a: Brand

$$\hat{p}^* = .07$$

$$p^* = .12$$

$$Z = -2.7$$

	MAXWELL HOUSE	TASTERS CHOICE	SANKA	BRIM	FOLGERS	NESCAFE
Brand Structure:	$\hat{P}(\underline{s}) = .07$	$\hat{P}(\underline{s}) = .05$	$\hat{P}(\underline{s}) = .08$	$\hat{P}(\underline{s}) = .09$	$\hat{P}(\underline{s}) = .04$	$\hat{P}(\underline{s}) = .06$
No structure:	$P(\underline{s}) = .18$	$P(\underline{s}) = .11$	$P(\underline{s}) = .12$	$P(\underline{s}) = .04$	$P(\underline{s}) = .02$	$P(\underline{s}) = .04$
	$Z = -2.9$	$Z = -1.3$	$Z = -.8$	$Z = 1.3$	$Z = .9$	$Z = -.5$

ALTERNATIVE b: Product Attribute

$$\hat{p}^* = .32$$

$$p^* = .11$$

$$Z = 11.3$$

GROUND CAFFEINATED	GROUND DECAFFEINATED	INSTANT CAFFEINATED REGULAR	INSTANT CAFFEINATED FREEZE DRIED	INSTANT DECAFFEINATED REGULAR	INSTANT DECAFFEINATED FREEZE DRIED
$\hat{P}(\underline{s}) = .47$	$\hat{P}(\underline{s}) = .19$	$\hat{P}(\underline{s}) = .37$	$\hat{P}(\underline{s}) = .23$	$\hat{P}(\underline{s}) = .28$	$\hat{P}(\underline{s}) = .25$
$P(\underline{s}) = .11$	$P(\underline{s}) = .03$	$P(\underline{s}) = .13$	$P(\underline{s}) = .09$	$P(\underline{s}) = .09$	$P(\underline{s}) = .13$
$Z = 8.4$	$Z = 4.1$	$Z = 5.5$	$Z = 3.3$	$Z = 4.1$	$Z = 2.4$

FIGURE 3: HYPOTHESIS TESTING FOR THE COFFEE MARKET

ALTERNATIVE c: Use

$$\begin{aligned}\hat{p}^* &= .40 \\ p^* &= .46 \\ Z &= -1.5\end{aligned}$$

	Morning (A.M.)	Afternoon & Evening (P.M.)
Use Structure:	$\hat{P}(\underline{s}) = .40$	$\hat{P}(\underline{s}) = .41$
No structure:	$P(\underline{s}) = .44$	$P(\underline{s}) = .48$
	$Z = -1.5$	$Z = -2.8$

ALTERNATIVE d: Users

$$\begin{aligned}\hat{p}^* &= .28 \\ p^* &= .59 \\ Z &= -8.69\end{aligned}$$

	Heavy	Light
User Structure:	$\hat{P}(\underline{s}) = .34$	$\hat{P}(\underline{s}) = .11$
No Structure:	$P(\underline{s}) = .75$	$P(\underline{s}) = .18$
	$Z = -9.3$	$Z = -1.2$

FIGURE 3: HYPOTHESIS TESTING FOR THE COFFEE MARKET (Cont.)



Before we select a structure for managerial use, the data are further analyzed to explore the effects of compound structure, the use of probabilities based on consideration sets, and the application of other market structure methods.

### Exploratory Testing of Compound Structure

We explored compound structure by calculating individual probabilities for various submarket structures. The assignment of individuals to the best fitting competitive structure led to the identification of some differences in consumers' views of the competition. Table 8 shows the number of people who (heuristically) fit one structure better than others.

TABLE 8: VARIATION IN COMPETITIVE STRUCTURE BY CONSUMER

<u>Dominant Structure</u>	<u>Number in Which One Structure Dominates</u>	<u><math>\hat{P}^*</math></u>
1. Ground/Instant	84	.94
2. Caffeinated/Decaffeinated	53	.94
3. Brand	1	.85
4. No dominant two-way structure	96	

The values of  $\hat{P}^*$  are high for groups described by ground/instant and caffeinated/decaffeinated structures. Based on our heuristic classification, the compound structure fits better than in the homogeneous case which has a  $\hat{P}^* = .73$  for an overall ground/instant structure and  $\hat{P}^* = .64$  for an caffeinated/decaffeinated structure. This is, of course, subject to confirmatory testing since the heuristic procedure is an attempt to maximize  $\hat{P}^*$ . These two groups may represent significant heterogeneity while the remaining 60 percent of the sample are probably not better described by one sub-structure than others. (Eight-four people were equally well described by

Alternatives 1 and 2, six people by Alternatives 2 and 3, and six people by Alternatives 1, 2, and 3.) Although some compound structure is evident, most compound structure seems to be simplified versions of alternative b. Thus, we choose the overall attribute structure shown in figure 3b is an adequate managerial summary description of the market.

Consideration Sets

In our method competitive structure can derive significance from consumer preferences and/or the compositions of consideration sets. In order to get an indication of the relative magnitude of these two effects we re-estimated the individual product deletion probabilities assuming equal preference across the products in each consumer consideration set after the most preferred had been removed. See Equation (18), hybrid technique. These values are shown in Table 9 for the six submarkets.

TABLE 9: PROBABILITIES BASED ON  
CONSIDERATION SETS -  $\hat{P}(s)$

	WITH PREFERENCE	CONSIDERATION SET ONLY
GROUND/CAFFEINATED	.47	.44
GROUND/DECAFFEINATED	.19	.18
INSTANT/CAFFEINATED/REGULAR	.37	.38
INSTANT/CAFFEINATED/FREEZE DRIED	.23	.16
INSTANT/DECAFFEINATED/REGULAR	.28	.27
INSTANT/DECAFFEINATED/FREEZE DRIED	.25	.22
OVERALL	.32	.30

They are very similar to preference based probabilities and the overall significance level drops only slightly ( $Z = 11.3$  to  $Z = 10.1$ ). In this case the consideration set contains much of the competition structure information.

### Comparison to Other Methods

The final phase of exploratory analysis of the data was the examination of the results of applying two alternative methods. Kalwani (1979) has applied hierarchical clustering procedures to our forced switching data. He found the best definition to be ground versus instant coffee with a second level division into caffeinated and decaffeinated variants. He tested several measures of proximity based on switching: (1) "last purchase" from "purchase previous to last purchase," (2) "last purchase" to "next purchase" (obtained in call-back), and (3) forced switching in laboratory. In all clustering, the cophenetic correlations were high (.814 to .887). No statistical basis for determining the best partitioning was available, but Kalwani felt that the results based on forced switching were superior since they tended to cluster products together that shared a common attribute and agreed with managerial priors. His results are consistent with the findings reported in our hypothesis testing. Kalwani also attempted to apply the Hendry model to the data. First he found the switching matrix did not meet the required equilibrium assumption. The observed switching was much below the theoretical switching (overall theoretical  $k_w = .53$ , empirical  $k_w = .28$ ). Next he calculated values of the switching constant  $k_w$  for possible partitions. The reported values of the theoretical  $k_w$ 's were very similar for all alternatives (.48 to .53) and uniformly above the observed values (.12 to .28). He could not reach any conclusion on the hierarchical form based on the Hendry methodology applied to this data. Although the Hendry method was not informative here, it does not reject an attribute-based partitioning and Kalwani's clustering analysis is consistent with the product based on submarkets shown in Figure 3b.

### Managerial Analysis

"Best" Structure: The first task is to define the best structure for competition. The formal hypothesis testing resulted in only one hypothesis that could explain the

observed switching significantly better than the null hypothesis of no competitive structure. The exploratory analysis indicates compound structure is not likely to be a serious managerial concern and that the analysis of the data by a clustering methodology is consistent with the product attribute analysis. An analysis of the forced choice shopping data collected in the laboratory store allows a test of convergent validation of the measures used to test the product attribute hypothesis (see Figure 3b). In Table 10 the probabilities from the preference analysis, shopping measures, the pooled probabilities, and the "no structure" probabilities are shown. The shopping probabilities are substantially greater than the "no structure" values and the product attribute structure is significantly better than the null model based on both the shopping ( $Z = 9.4$ ) and pooled probabilities ( $Z = 14.8$ ). Based on the pooled data for the attribute submarkets, we examined the significance of an attribute structure for each brand where at least 10 consumers preferred most that brand (see Table 11). In nine of the ten cases, the tests indicate significance at least at the 10 percent level. Thus, from the perspective of the significant brands, we select the best description of the competitive structure of the coffee market based on product attributes for managerial use in determining marketing strategy.

TABLE 10: POOLING OF LABORATORY SHOPPING FORCED SWITCHING AND  
INDIVIDUAL LOGIT-BASED PROBABILITIES

<u>SUBMARKET</u>	$\hat{P}(s)$			NO STRUCTURE ( $\hat{P}(s)$ )	Z-STATISTIC POOLED DATA VERSUS NO STRUCTURE
	<u>INDIVIDUAL</u>	<u>SHOPPING</u>	<u>POOLED</u>		
GROUND/CAFFEINEATED	.47	.47	.47	.11	11.4
GROUND/DECAFFEINATED	.19	.20	.20	.03	5.1
INSTANT/CAFFEINATED/REGULAR	.37	.28	.33	.13	6.2
INSTANT/CAFFEINATED/FREEZE DRIED	.23	.25	.24	.09	4.9
INSTANT/DECAFFEINATED/REGULAR	.28	.18	.24	.09	4.2
INSTANT/DECAFFEINATED/FREEZE DRIED	.25	.32	.28	.13	4.1
<u>AGGREGATE Z STATISTIC</u>	11.3	9.4	14.8		

TABLE 11: BRAND LEVEL SIGNIFICANCE  
TESTS FOR ATTRIBUTE HIERARCHY\*

	$\hat{P}_i(\underline{s})$	$P_i(\underline{s})$	Z
GROUND/CAFFEINATED			
MAXWELL HOUSE	.49	.10	11.4
CHOCK FULL O'NUTS	.60	.14	5.3
INSTANT/CAFFEINATED/REGULAR			
INSTANT MAXWELL HOUSE	.33	.10	6.5
NESCAFE	.37	.19	2.1
INSTANT/CAFFEINATED/FREEZE DRIED			
MAXIM	.30	.14	2.5
TASTERS CHOICE	.21	.06	4.5
INSTANT/DECAFFEINATED/REGULAR			
NESCAFE	.18	.16	.2
SANKA	.25	.08	4.9
INSTANT/DECAFFEINATED/FREEZE DRIED			
SANKA	.31	.14	2.3
BRIM	.35	.14	2.8

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\*Table includes only brands where more than 10 consumers preferred most that brand.

Managerial Implications: The appropriate product strategy for a company depends, in part, upon which products the firm now offers. Table 12 shows the existing brands (1977) of three major manufacturers. If Nestle' is considered, there is a gap in the product line coverage since they now offer no ground coffees. Our analysis suggests ground coffee is a separate market from instant. Substitution between a Nestle' ground coffee and Nestle' brands of instant coffee would be low based on this definition of competition. Our tests suggest that Nestle' should reconsider its coverage of submarkets. In 1977 Nestle' had a high share in the instant submarkets, but a zero market share in the ground coffee submarket. For strategic analysis, Nestle' might wish to calculate share with respect to the instant submarket rather than with respect to the entire coffee market.

The development of a new ground coffee may be a major opportunity for Nestle'. Its proven ability to market instant coffees suggests it has the capability to advertise, promote, and distribute a ground coffee. This opportunity is suggested by our analysis and should be subjected to further analysis. Perceptual maps of the ground, caffeinated and decaffeinated submarkets could be drawn to more specifically define positioning opportunities and the potential of a new entry can be estimated. See examples in Urban, Johnson and Brudnick, 1981 and Urban, Carter and Mucha, 1983. These opportunity identification activities are useful first steps in the development of a new product offering. For more details on how to balance these considerations with sales potential, penetration, scale, input, reward, risk, and match to the organization's capabilities, see Urban and Hauser (1980, Chapter 5).

TABLE 12

## BRAND OFFERINGS BY SELECTED FIRMS (1977)

	Ground		Instant			
	Caffeinated	Decaffeinated	Caffeinated		Decaffeinated	
			Regular	Freeze-Dried	Regular	Freeze-Dried
Nestle	No Brand	No Brand	Nescafe	Tasters' Choice	Nescafe	Tasters' Choice
General Foods	Maxwell House Yuban	Sanka Brim	Maxwell House Yuban	Maxim	Sanka	Sanka Brim
Procter & Gamble	Folger's No Brand		Folger's	No Brand	(High Point)	No Brand

( ) = in test market at time of study



General Foods has a different strategic problem. Each submarket is covered -- General Foods already has many coffee offerings. If Maxwell House and Yuban (Sanka and Brim) are not well differentiated within their submarkets, then General Foods may have too many offerings and should consider consolidating its brands. It could also direct its product development efforts at creating a new submarket. The methodology we propose only describes the current market situation. A new product can fit into a submarket or, possibly, create a new submarket. For example, perhaps General Foods could develop fully brewed coffee that is frozen in plastic cups and can be heated in a microwave oven. This may add a new submarket to the structure. This is an opportunity that could be pursued by product development and testing. As the idea is converted into a concept, product, and complete design, data would be collected to confirm the strategic opportunity and balance it against other considerations.

Proctor and Gamble has recently entered the coffee market with the Folgers brand. Their strategy appears to be to build their business by sequentially entering each submarket in order of the sales potential of the submarkets. High Point is now in the national market. A next entry for them to evaluate could be a freeze dried instant. The opportunity of a ground decaffeinated coffee could also be evaluated.

#### OTHER APPLICATIONS

The coffee example shows how the testing procedure can be used to define the basis for market share, identify opportunities for new products, and provide a structure to evaluate product strategy. We therefore call the model, measurement, and testing system PRODEGY (Product Strategy). It has been applied seven times in the past and other applications are now in process.

In one case on home cleaning products alternative competitive market hypotheses included forms of the products (e.g., spray, liquid, powder, foam, aerosal), user locations (e.g., kitchen, bathroom), user tasks (e.g., glass, counters, floor, metal chrome), user intensity (e.g., light cleaning, heavy cleaning), and individual attributes (e.g., "sophisticated" cleaners who use many special purpose products versus

"basic" cleaners who use a few generic products). The selected submarket structure was a "use" partitioning. It reflected a new perspective on the market and indicated a major gap in the firm's product line. Previously, four attempts had been made to develop a new brand in this product class, but all failed based on pre-test market analysis. The new opportunity identified by PRODEGY was pursued and it succeeded in pre-test market analysis. It is now being successfully test marketed and the brand manager credits the competitive market structure as a major contributor to the product's success.

In an application to the beer market, brand, product form, and usage level structures were evaluated. The best partitioning was found to be consistent with the firm's previous beliefs and reflected an empirical confirmation of its strategic assumption on competitive boundaries. This empirical confirmation of the firm's prior views increased confidence in their belief that they were adequately covering the existing market for beer. They turned their development attention to creating a new partition in the market by searching for a new form of the product (e.g., super light beer).

In a study of detergents, a new competitive brand recently introduced in the market was placed in the structure. PRODEGY indicated to the firm that a pourable powder detergent was not creating a new subsection in the competitive structure. The new product was perceived as competing in an existing product form partition. The firm responded by repositioning their existing brand rather than spending large sums of money to introduce a major brand against the new entrant.

In the final consumer goods case, food products were studied. It was found that the "pre-packaged" and "deli" brands reflected separate submarkets in the category. The firm previously had refrained from selling a "deli" product because they felt it would compete with their line of "prepackaged" goods. After the study indicated this was not true, they began development of an entrant into the deli market with the prospect of major sales increases. Perceptual maps indicated the firm had the product strength to compete in the deli segment and concept testing was begun to test the transferability of the firm's brand name from the pre-packaged to deli submarkets.

Two applications have been made to industrial products. In a study of financial decision support system software for a firm selling a sophisticated software product, the "no structure" hypothesis could not be rejected and the firm's strategic space was described by a single perceptual map (Burrow and Burns, 1982 and Borshberg and Elkins, 1983). Financial planners viewed the major dimensions as "power" (math capability, large data bases, financial functions and consolidation capabilities), "ease of use" (manipulate data and develop models easily, understandable documentation, and easy to learn), and "vendor quality" (reputation and support). Planners considered both simple programs such as Visicalc or Supercalc and complex programs such as Express or XSIM for each specific use (e.g., long range planning, budgeting, and financial reporting). A map was drawn by graphing the attribute levels per dollar. This suggested that complex products would not be successful unless they had both good power and ease of use per dollar. The firm commenced a program to improve its ease of use per dollar by developing menu driven products which used subsets of the full software systems capabilities and reduced its prices. It also committed to monitoring the market to see if partitions would develop as users became more knowledgeable and their problem solving needs increased.

Another industrial application was to heart pacemakers, (Kasinkas, 1982). In this application the customer was the doctor. For each of three patient symptomologies doctors indicated which brands and types of pacemakers they would consider (a list of 40 products was supplied) and their probability of implanting that device. Direct estimates of probabilities of purchase (Juster, 1966, and Morrison, 1979) and the assumption of equal probabilities over the consideration set were used to estimate the product deletion probabilities. The use of laboratory forced switching was not feasible in this study. The pacemaker study found submarkets defined by product attributes (e.g. programmable versus non-programmable units) as a significantly better description of the market than "no structure." The implication for the manufacturers was not to drop its older simple non-programmable units because a submarket persists for it and to continue to develop more advanced units to capture the sophisticated submarket.

## SUMMARY AND RESEARCH ISSUES

This paper has presented a model and measurement methodology to estimate the structure of competition. In seven applications it produced encouraging statistical significance and managerial insight; however, several issues deserve further research.

One definition has been proposed in this paper for the description of competitive structure of a market. A statistical test was derived based on this definition and several procedures were suggested to collect data for the tests. We provide empirical applications based on our methods, but it would be useful to conduct comprehensive comparative empirical studies across additional methods (e.g., overlapping clustering) and alternative data sources (e.g., panel and UPC data). This is a subject for future research.

A technical issue to be resolved in the model is what to do if only one product defines a competitive submarket. In this case, the model criterion based on forced switching is ambiguous. Although a single product submarket cannot be tested formally, intuitively one would be indicated if consumers would consider that product as the only acceptable alternative for a specific use. In terms of our measures, a consideration set of one brand and a large proportion of respondents refusing to buy in the laboratory store would indicate this condition. Research may improve procedures for identifying and testing single product branches, but in practice, we rarely observe such submarkets because competition usually develops quickly if the initial product is successful.

The empirical success of using consideration sets to estimate individual probabilities opens the possibility of monitoring the competitive structure over time through low cost telephone surveys or UPC panels. The ability to represent the dynamics of the competitive relationships would be very useful in strategy formulation. It could allow low cost measurement of the emergence of new submarkets.

New applications are underway in several markets to further assess the empirical adequacy of the model and its managerial relevance. Our early applications suggest that the proposed model could be a useful tool to aid in market strategy formulation.

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APPENDIX

USE AND BRAND CONSIDERATION BY OCCASION (COFFEES)

In interviews with 295 coffee drinkers (greater than one cup per day), 808 uses were evoked across six major use classes (average is 2.7 uses per person). The table shows the proportion of the respondents who evoked each use and the proportion who would consider a given brand for those who evoked given use. For example, 8.5 percent of the respondents who evoked breakfast as a use would consider Brim (Instant) as a product for this use.

	<u>Breakfast</u>	<u>Day Alone</u>	<u>Day Others Present</u>	<u>Lunch</u>	<u>Supper</u>	<u>Evening</u>
Percent Evoking Use	96.3	41.7	23.4	30.2	43.7	43.1
<u>Percent Who Would Consider Brand:</u>						
Brim (Instant)	8.5	6.5	12.1	7.9	8.4	7.1
Folgers (Instant)	6.3	6.5	3.0	6.7	6.7	6.3
Folgers (Ground)	7.7	9.8	9.1	3.4	5.9	7.9
Maxwell House (Instant)	32.0	26.0	40.9	36.0	33.6	27.6
Nescafe	10.6	11.4	18.2	9.0	8.4	15.0
Nescafe (Decaf)	10.6	8.9	6.1	12.4	10.1	7.1
Maxim	12.0	15.4	7.6	13.5	13.4	11.8
Taster's Choice	20.1	18.7	24.2	20.2	18.5	22.0
Taster's Choice (Decaf)	13.0	17.1	16.7	20.2	17.6	17.3
Sanka (Instant)	22.9	22.0	18.2	21.3	20.2	24.4
Sanka (Freeze Dried)	7.7	8.1	9.1	12.4	11.8	3.9
Sanka (Ground)	6.3	7.3	12.1	5.6	9.2	4.7
Chock Full O'Nuts	10.6	8.9	10.6	6.7	8.4	9.4
Hills Brothers	7.0	4.1	1.5	3.4	6.7	1.6
Maxwell House (Ground)	28.9	32.5	24.2	29.2	32.8	26.0

ADDENDUM TO: G.L. Urban, P.L. Johnson, and J.R. Hauser, "Testing Competitive Market Structure: A Methodology Based on Product Deletion Probabilities," Marketing Science, Vol. 3, 1984.

TECHNICAL ANALYSIS OF POTENTIAL VIOLATION OF THE ASSUMPTION OF  
INDEPENDENCE OF IRRELEVANT ALTERNATIVES IN THE LOGIT MODEL

The rank order preference method assumes the weak form of independence of irrelevant alternatives. McFadden, Train and Tye (1977) have developed a residual test procedure for the logit model. We applied it here to the probabilities from the rank order preference model. This table shows the chi-squared statistic associated with this test for the major products. In three cases (Maxwell House Ground, Maxim, and Taster's Choice caffeinated) violation of independence is indicated. In the other five cases, violation is not indicated.

Table A-1: Test of Independence of Independent Alternatives

<u>Product</u>	<u><math>\chi^2(df=9)</math></u>
Maxwell House Ground	17.1*
Taster's Choice (decaffeinated)	5.6
Nescafe	11.7
Maxim	17.2*
Instant Maxwell House	3.3
Chock Full O'Nuts (Ground)	6.4
Taster's Choice (caffeinated)	18.8*
Sanka (Instant)	11.0

\* = significant at the 10 percent level.

These results are mixed, but indicate the independence of irrelevant alternative assumption may be violated.

In order to determine the sensitivity of our probability estimation to this potential violation, we re-estimated the individual choice probabilities ( $P_C(j)$ ) and the forced switching probabilities ( $P_f(s)$ ) by hierarchical logit procedures (McFadden, 1980).

The use of the hierarchical logit model requires that we assume our preference measures (7 point) are interval scales. To investigate the effect of an interval scaling assumption, we compare the rank order preference and standard multinomial logit model in which the interval property is also necessary. The correlation of these choice probabilities to those estimated by the rank order model was high ( $\rho = .965$ ), so the probability estimates do not appear to vary much when the interval scaling property is assumed.

The correlation of the probabilities from the hierarchical logit to those from the multinomial logit model in the ground/instant structure was .84. This is not as high as the previous correlation and use of the probabilities from the hierarchical logit led to different forced switching probabilities ( $\hat{P}(s)$ ). For example,  $\hat{P}(s)$  for ground coffee was .42 with the hierarchical logit probabilities -- less than the .60 obtained with the multinomial Logit. One possible explanation of the difference was compound structure (see Table 10). In order to investigate this effect the hierarchical and multinomial logit probability estimates were obtained within the two separate groups of consumers -- ground/instant (84 observations) and caffeinated/decaffeinated (53 observations). Despite the small samples, within these groups the correlation returned to high levels ( $\rho = .95$  for the caffeinated/decaffeinated group) and the aggregate probabilities were high ground/instant group and  $\rho = .91$  for the ( $\hat{P}^* = .84$  for the ground/instant group and  $\hat{P}^* = .75$  for caffeinated/decaffeinated group). These findings suggest extreme caution should be exercised in applying hierarchical logit procedures when compound structure is possible.

Once compound structure is considered, the sensitivity to possible violation of the assumption of independence of irrelevant alternatives is low since the hierarchical and non-hierarchical logit procedures yield very similar individual probability estimates ( $P_c(j)$ ).

The absence of an empirical problem with the independence of irrelevant alternatives can be explained in part by the research design which measures preferences only over individual consideration sets. The consideration set is the consumer's self-screening of the alternatives and the designation of those "relevant" to his or her choice. The empirical analysis here is consistent with the notion that a CRM model describes individual choices across the consideration set and that the competitive structure is primarily the result of an aggregate competitive structure in the market rather than the sum of individual hierarchical choice models. This is consistent with the empirical analyses described in the text.