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Search for Invisible Decays of the $Y(1S)$

We search for invisible decays of the $Y(1S)$ meson using a sample of $91.4 \times 10^6$ $Y(3S)$ mesons collected at the BABAR/PEP-II $B$ factory. We select events containing the decay $Y(3S) \rightarrow \pi^+ \pi^- Y(1S)$. 

The nature of dark matter is one of the most challenging issues facing physics. Observation of standard model (SM) particles coupling to undetectable (invisible) final states might provide information on candidate dark matter constituents. In the SM, invisible decays of the Y(1S) meson proceed by $b\bar{b}$ annihilation into a $\nu\bar{\nu}$ pair, with a branching fraction $B(Y(1S) \rightarrow \text{invisible}) \approx 1 \times 10^{-3}$ [1], well below the current experimental sensitivity. However, the dominant background candidates could couple weakly to SM particles to enhance the invisible branching fraction to well below the current experimental sensitivity. However, it is primarily used to correct the peaking background for imprecisely known branching fractions. The “three-track” subsample is composed of events containing two pions and only one high-momentum track, consistent with two-body decay of the Y(1S) where both final-state particles are detected. It is used to adjust for inaccuracies in the modeling of track acceptance.

We select events in the invisible subsample by requiring that there are exactly two tracks originating from the interaction point (“IP tracks”) with opposite electric charge. An IP track is required to have a point of closest approach to the interaction point within 1.5 cm in the plane transverse to the beams and within 2.5 cm along the z axis. We further require these tracks to each have c.m.-frame momentum $p < 0.8$ GeV/$c$, consistent with pions from the dipion transition. The dipion system is required to have an invariant mass satisfying $M_{\pi\pi} \in [0.25, 0.95]$ GeV/$c^2$, compatible with kinematic boundaries ($M_{\pi\pi} \in [2M_{\pi}, (M_{Y(3S)} - M_{Y(1S)})]$) after allowing for resolution effects. The dipion recoil mass is

$$M_{\text{rec}}^2 = s + M_{\pi\pi}^2 - 2\sqrt{s}E_{\pi\pi},$$

(1)

where $E_{\pi\pi}$ is the c.m. energy of the dipion system and $\sqrt{s} = 10.3552$ GeV/$c^2$. We require that $M_{\text{rec}} \in [9.41, 9.52]$ GeV/$c^2$. The efficiency of this selection for signal events is about 64%, due to the requirement of reconstructing the two pions. All selection criteria were finalized without looking at data in a narrower $M_{\text{rec}}$ “signal region” that, according to simulation, contains more than 99% of the signal (see discussion of signal shape below for the precise signal-region definition).

We select three-track and four-track events using the same dipion selection as in the invisible subsample. We search for high-momentum tracks from the Y(1S) decay [i.e., from $Y(1S) \rightarrow e^+ e^-$ or $Y(1S) \rightarrow \mu^+ \mu^-$]. We require that there be one or two additional IP tracks, each with $p^T > 2.0$ GeV/$c$. If either of these tracks passes electron-identification criteria, both are treated as electrons; otherwise, both are treated as muons. In the former case, we account for possible radiative energy loss due to
bremstrahlung by pairing an electron with a photon emitted close in angle and increasing the electron’s energy and momentum by the energy of this photon. When two high-momentum tracks are present, we require that they have opposite charge and a two-track invariant mass \( \in [9.00, 9.80] \text{GeV}/c^2 \). We remove photon conversions from these events by rejecting the event if either pion satisfies electron-identification criteria. This introduces a negligible efficiency loss: the probability of a pion to be misidentified as an electron is \( \approx 0.1% \). Finally, we require that the mass difference between the \( \pi^+ \pi^- \ell^+ \ell^- \) and \( \ell^+ \ell^- \) systems \( \in [0.89, 0.92] \text{GeV}/c^2 \).

At this stage, the background level in the invisible subsample is several orders of magnitude larger than any hypothetical signal. We reject most of this remaining background with a multivariate analysis (MVA), implemented as a random forest of decision trees [10]. The random forest algorithm is trained on signal-MC events and 5.3% of data outside of the signal region in \( M_{\text{rec}} \). The contribution of peaking components to these data is negligible. The data and signal-MC events used to train the MVA are excluded from the rest of the analysis, leaving 91.4 \( \times 10^6 \) \( Y(3S) \) events for use in the final result.

We use the following variables, which have been determined to be only weakly correlated with \( M_{\text{rec}} \), as inputs to the MVA: (1) the probability that the pions originate from a common vertex, (2) the laboratory polar angle and transverse momentum of the dipion system, (3) the total number of charged tracks, IP tracks or otherwise, reconstructed in the event, (4) booleans that indicate whether either pion has passed electron, kaon, or muon identification criteria, (5) the cosine of the angle (in the c.m. frame) between the highest-energy photon (\( \gamma_1 \)) and the normal to the decay plane of the dipion system, (6) the energy in the laboratory frame of the \( \gamma_1 \), (7) the total neutral energy in the c.m. frame, and (8) the multiplicity of \( K^0_{\text{S}} \) candidates, defined using the shape and magnitude of the shower resulting from interactions in the calorimeter.

The selection on the MVA output is optimized by choosing the threshold that achieves the minimum statistical uncertainty (dominated by background) on \( \mathcal{B}(Y(1S) \rightarrow \text{invisible}) \) and, in the null signal hypothesis, the lowest upper limit at the 90% C.L. Both were achieved by requiring an MVA output >0.8 (where the full range is 0 to 1). The efficiency of this criterion for signal-MC events is 37.0%, as compared to 0.8% for data events outside of the signal region. The total efficiency of all trigger and event selection requirements is determined from signal-MC simulation to be 16.4%.

Figure 1 shows the resulting \( M_{\text{rec}} \) distribution for events in the invisible subsample. We extract the peaking yield by an extended unbinned maximum likelihood fit, with the nonpeaking background described by a first-order polynomial. The signal and peaking background should have the same shape in \( M_{\text{rec}} \). We describe this shape by a modified Gaussian function with a common peak position (\( \mu_0 \)), independent left and right widths (\( \sigma_{L,R} \)), and non-Gaussian tails (governed by parameters \( \alpha_{L,R} \)). The functional form on either side of the peak is

\[
f_{L,R}(M_{\text{rec}}) = \exp\left[-(M_{\text{rec}} - \mu_0)^2/(2\sigma_{L,R}^2)\right] + \alpha_{L,R}(M_{\text{rec}} - \mu_0)^2.
\]

We determine the parameters of this probability density function (PDF) by fitting \( M_{\text{rec}} \) in the four-track data subsample. The signal region, excluded when training the MVA, is defined as the region in Fig. 1, which is \( <5\sigma_{L,R} \) from the peak position, \( M_{\text{rec}} \in [9.4487, 9.4765] \text{GeV}/c^2 \).

The fit to the invisible subsample then determines all of the parameters of the nonpeaking background PDF, the yield of the nonpeaking background, and the yield of the peaking component. Using an MC-based method, we find that the fit accurately returns the peaking contribution over a wide range of input values. The result for the peaking yield in data is \( 2326 \pm 105 \) events.

Using a second-order polynomial for the nonpeaking background results in no change in the extracted peaking yield. The systematic uncertainty on that yield associated with the fixed parameters in the signal PDF is estimated by varying those parameters in the fit. We find an uncertainty of 18 events.

We next estimate the contribution of background to the peak. The MC simulation predicts 1019 \( Y(1S) \rightarrow e^+e^- \) events, 1007 \( Y(1S) \rightarrow \mu^+\mu^- \) events, 92 \( Y(1S) \rightarrow \tau^+\tau^- \) events, and \( 2.9 \pm 1.3 \) \( Y(1S) \rightarrow \text{hadrons} \) events. These predictions depend upon branching fractions which have significant uncertainties [5] and on the accuracy of the modeling of event reconstruction and selection. We use four-track and three-track data and MC subamples to test and correct the MC prediction of 2122 total events.
We first use the four-track subsamples to calibrate the product of the branching fractions for \(Y(1S) \rightarrow \ell^+ \ell^-\) and \(Y(3S) \rightarrow \pi^+ \pi^- Y(1S)\) and the dipion reconstruction efficiency. We compare the event yields between four-track data and MC subsamples when the positively charged lepton is emitted in the central section of the tracking system, \(|\cos(\theta_{rec})| < 0.3\) (laboratory-frame angle). The simulation underestimates the number of events in data by a factor of \((1.088 \pm 0.012)\). This is plausible in light of the branching fraction uncertainties \([4.7\%\text{ on the dipion transition, } 2.5\%\text{ on the } Y(1S) \text{ decay} \]) and track reconstruction uncertainties \((=0.5\%\text{ per track})\). Since the effect of the high-momentum track reconstruction has a negligible contribution here, this data or MC correction factor is applied to all of our MC-simulation subsamples. For the four-track subsample, Fig. 2(a) shows that the distribution of the high-momentum tracks in the detector is very well described by the MC simulation at all polar angles.

We next compare the data and MC efficiencies for reconstructing the single lepton in the three-track subsample. Any discrepancy would imply a complementary mistake in the invisible peaking background. Given the c.m.-frame polar angle coverage of the detector, for three-track events the high-momentum lepton in the forward direction often escapes detection and thus the detected lepton is in the backward direction. We compare the MC and data laboratory-frame polar angle distributions for these events in Fig. 2(b). The three-track subsample, in contrast to the four-track subsample, has a significant nonpeaking background. Hence three-track peaking yields versus polar angle are determined by using the \(M_{\text{rec}}\) fit described above and applying an event-weighting technique \([11]\). The MC simulation describes the distribution well everywhere except at \(\cos(\theta_{rec}) < -0.84\), where the simulation overestimates the reconstruction rate.

For leptons in this far-backward region, we use the ratio of data to simulation versus lepton \(\cos(\theta)\) from Fig. 2(b) as the basis of an accept-reject method applied to the high-momentum track. When this method removes the track, it in effect reassigned a three-track event to the invisible category. We also weight the reassigned events by the ratio of simulated trigger efficiencies for the three-track and invisible subsamples and assign 100% uncertainty to this difference in trigger efficiency. Applying this additive correction after the scaling correction (from the four-track subsample), the total peaking background estimate increases from 2122 events to \((2451 \pm 38)\) events.

We test the prediction of the contribution of nonleptonic \(Y(1S)\) decays to the peaking background using an additional control sample. Events in this sample contain only two tracks (the pions) and pass all other criteria for the invisible subsample, except that the MVA requirement is replaced by a requirement that the \(\gamma_1\) has energy \(>0.250\text{ GeV}\). This selects a set of events which is almost disjoint from the invisible subsample, since the MVA \(>0.8\) requirement results in a steep falloff in efficiency versus \(\gamma_1\) energy near \(0.250\text{ GeV}\). We compare this energy distribution in data (using the weighting technique \([11]\)) to that from simulation. As the \(\gamma_1\) energy approaches \(0.250\text{ GeV}\) from above, we find that the MC simulation underestimates the data by no more than a factor of 4. Since the expected contribution of these events to the peaking background is \(0.14\%\) of the total, we assign \(0.6\%\) \((15\text{ events})\) as an additional systematic uncertainty on the peaking background for a total of \(\pm 41\) events.

A number of multiplicative systematic corrections and uncertainties to the peaking background also enter, in a fully correlated manner, when the extracted signal yield is converted to the \(Y(1S) \rightarrow \text{invisible}\) branching fraction. The first such contribution is the \(1.088 \pm 0.012\) correction factor derived from the four-track subsample. But this does not account for trigger and MVA effects which might differ for the invisible and four-track subsamples. Since events used to train the MVA have already passed the trigger requirements, we first study the effect of trigger selection on data. The \(\text{BABAR}\) trigger consists of a hardware and a software stage. The latter is tested by using a heavily prescaled sample of events which bypassed it. We apply the software-level trigger to these events and find that the ratio of efficiencies in data and MC simulation is

![Graphs](image-url)

**FIG. 2** (color online). The distribution of \(\cos(\theta)\), where \(\theta\) is the laboratory-frame polar angle of (a) the positively charged high-momentum track in the four-track subsample and (b) the single high-momentum track in the three-track subsample. The normalization correction from the four-track subsample has been applied to the MC yields in both cases.
0.997 ± 0.009. This ratio is taken as a correction to the signal efficiency and the peaking background. To assess how well the impact of the hardware trigger on the two pions is simulated, four-track events are used, since their trigger decision is based largely on the two high-momentum lepton tracks. We apply to the pions a set of selection criteria which approximate those applied by the hardware trigger. The data and MC efficiencies for these requirements differ by 2.2%. Since this test is done on samples for which the hardware trigger is only approximated, we take this difference as a systematic uncertainty rather than apply a correction for it.

After applying the approximate hardware trigger criteria to the four-track subsamples for both \( Y(3S) \) MC simulation and data, we apply the nominal MVA selection to both. The relative difference in efficiency between these MC and data subsamples is 4.0%. Since the hardware trigger is again only approximated for this test, we apply no correction for the difference but assign it as a systematic uncertainty on the MVA selection.

Adding the multiplicative uncertainties in quadrature, the total correlated systematic uncertainty is 4.8%. The final corrected prediction for the peaking background is (2444 ± 123) events, including the uncorrelated uncertainty of 41 events. From this we determine the signal yield to be (−118 ± 105 ± 124) events, where the errors are statistical and systematic, respectively. To obtain \( B(Y(1S) \rightarrow \text{invisible}) \), we divide this by the signal efficiency, the number of \( Y(3S) \) mesons, the branching fraction for the dipion transition (4.48% [5]), and the correction factors (1.088 × 0.997). The factor derived from the four-track subsample includes a possible adjustment of \( B(Y(1S) \rightarrow e^+e^-) + B(Y(1S) \rightarrow \mu^+\mu^-) \), not relevant for signal. We take this adjustment to be 1.000 ± 0.025 [5] and remove it by assigning an additional systematic uncertainty of 2.5%. Taking correlations into account, we determine that \( B(Y(1S) \rightarrow \text{invisible}) = (−1.6 ± 1.4(\text{stat}) ± 1.6(\text{syst})) \times 10^{-4} \).

Lacking evidence for this decay, we use a Bayesian technique to set an upper limit on the branching fraction. We convolute the statistical likelihood, a function of \( B(Y(1S) \rightarrow \text{invisible}) \), with Gaussian functions representing the systematic error. We assume a prior probability that is flat in branching fraction and integrate the likelihood from 0 to a value such that 90% of the total integral above 0 is enclosed. The resulting limit is \( B(Y(1S) \rightarrow \text{invisible}) < 3.0 \times 10^{-4} \) at the 90% C.L.

In conclusion, we search for invisible decays of the \( Y(1S) \) meson. We do so by looking for evidence of the decay of the \( Y(1S) \) into undetectable final states recoiling against the dipion system in \( Y(3S) \rightarrow \pi^+\pi^-Y(1S) \), using

a sample of \( 9.14 \times 10^6 \) \( Y(3S) \) mesons. We find no evidence for \( Y(1S) \rightarrow \text{invisible} \) and set an upper limit on its branching fraction at \( 3.0 \times 10^{-4} \) at the 90% C.L. This limit is almost an order of magnitude closer to the SM prediction than the best previous limit.

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