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Citation

As Published
http://dx.doi.org/10.1109/tbme.2009.2039213

Publisher
Institute of Electrical and Electronics Engineers

Version
Final published version

Accessed
Wed Oct 17 01:48:53 EDT 2018

Citable Link
http://hdl.handle.net/1721.1/62018

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Using Point Process Models to Compare Neural Spiking Activity in the Subthalamic Nucleus of Parkinson’s Patients and a Healthy Primate

Sridевi V. Sarma∗, Member, IEEE, Uri T. Eden, Ming L. Cheng, Ziv M. Williams, Rollin Hu, Emad Eskandar, and Emery N. Brown, Fellow, IEEE

Abstract—Placement of deep brain stimulating electrodes in the subthalamic nucleus (STN) to treat Parkinson’s disease (PD) also allows the recording of single neuron spiking activity. Analyses of these unique data offer an important opportunity to better understand the pathophysiology of PD. Despite the point process nature of PD neural spiking activity, point process methods are rarely used to analyze these recordings. We develop a point process representation of PD neural spiking activity using a generalized linear model to describe long- and short-term temporal dependencies in the spiking activity of 28 STN neurons from seven PD patients and 35 neurons from one healthy primate (surrogate control) recorded, while the subjects executed a directed-hand movement task. The model used the point process model to characterize each neuron’s bursting, oscillatory, and directional tuning properties during key periods in the task trial. Relative to the control neurons, the PD neurons showed increased bursting, increased 10–30 Hz oscillations, and increased fluctuations in directional tuning. These features, which traditional methods failed to capture accurately, were efficiently summarized in a single model in the point process analysis of each neuron. The point process framework suggests a useful approach for developing quantitative neural correlates that may be related directly to the movement and behavioral disorders characteristic of PD.

Index Terms—Deep brain stimulation (DBS), Parkinson’s disease (PD), point processes, spike trains.

I. INTRODUCTION

The use of chronic deep brain stimulation (DBS) is now a well-established therapy for Parkinson’s disease (PD) [5], [23]. Prior to the placement of the stimulating electrodes in the commonly targeted subthalamic nucleus (STN), microelectrode recordings are routinely performed to help to locate the STN and to insure proper placement of the electrode. The need to record neural spiking activity to insure proper electrode placement gives neurophysiologists a unique opportunity to learn directly about the pathological properties of STN neurons in human patients with PD [1], [4], [5], [10], [26], [29].

To date, the analyses of these unique data typically use several different statistical techniques to characterize the spiking properties of the STN neurons [14]. Short-term history dependence within a spike train is analyzed with interspike interval histograms [37]. Long-term history dependence related to neural oscillations is often studied in the frequency domain using power spectra by transforming the spike train into a continuous-valued signal before computing its Fourier transform [22], [28]. Movement-related properties are determined from a tuning curve computed as the average spike rates across multiple trials in each movement direction [20], [21], [33].

More recently, point process methods have been used to analyze the spike train activity for a broad range of neural systems [6], [25], [30], [35], [42], [43]. Despite the point process nature of STN spiking activity, point process methods are not routinely used to analyze these recordings. Some studies have applied Poisson models to neuronal spiking data of PD patients, but these models fail to capture any temporal dependencies that exist in the spiking activity, such as refractoriness, bursting, and oscillations [26], [29]. Zelnikera et al. [40] analyzed the stochastic structure of the short- and long-term dependencies as two Poisson processes with different rate constants. Although informative, this approach restricted the summary of the temporal dependencies in the data to the differences between the two rate constants. Eden et al. recently demonstrated that a single point process model could be used to characterize the spiking activity in STN neurons recorded from PD patients executing a directed-hand movement task [18]. The model used a generalized linear model (GLM) [27] to represent the point process conditional intensity function (CIF) in terms of both short- and long-term history dependence [35]. The single model captured oscillations, bursting, directional tuning, and thus, obviated the need for multiple different analyses. In addition, the model identified a previously undescribed period of decreased spiking propensity 20–30 ms following a spike immediately prior to movement [18].

To use this recent work to understand more clearly the extent to which the features of the neural spiking activity observed in
the PD patient’s are signatures of this pathological condition, it
would be ideal to compare these features with similar features
derived from spiking activity recorded in healthy human subjects
executing the same movement task. Because this is not possible
for obvious ethical reasons, we decided to compare the STN
neural spiking activity of seven PD patients with STN activity
recorded from a healthy nonhuman primate, rhesus monkey,
executing the movement task. The nonhuman primate serves as
a surrogate control. We use GLMs to formulate point process
models [15], [34] to characterize the relative contribution of
intrinsic factors (e.g., short- and long-term history effects) and
extrinsic factors (e.g., the impact of movement direction) on
the probability that the neuron will spike at any given time
[35]. Once adequate goodness-of-fit is established, we use the
model parameters to analyze the relative importance of bursting,
oscillations, directional tuning in characterizing differences in
spiking propensity of neurons in the two groups.

II. METHODS

A. Human Subjects

Seven patients undergoing deep brain stimulator placement
for the treatment of PD were included in the study. All pa-
tients had idiopathic PD with a Hoehn–Yahr score [41] of three
or higher and had a documented response to the treatment
of PD were included in the study. All pa-

C. Primate Subject

One adult male rhesus monkey (macaca mulatta), a.k.a.
“Bohr,” was included in this study. A titanium head post, plastic
recording chamber, and scleral search coil were surgically im-
planted in accordance with guidelines set by the animal review
committee at Massachusetts General Hospital. Neuronal activity
was amplified, bandpass filtered between 200 Hz–5 kHz, and
sampled at 20 kHz. Spikes were sorted offline using a standard-
tized template-matching algorithm (Cambridge Electronics Design,

D. Behavioral Task

Once the microelectrodes were in the STN, the subjects
viewed a computer monitor and performed a behavioral task
by moving a joystick with the contralateral hand. The joystick
was mounted such that movements were in a horizontal orienta-
tion with the elbow flexed at approximately 45°. The behavioral
task began with the presentation of a small central fixation point.
After a 500 ms delay, four small gray targets appeared arrayed
in a circular fashion around the fixation point (up, right, down,
and left). After a 500–1000 ms delay, a randomly selected target
turned green (target cue (TC)) to indicate where the subject is
to move. Then, after another 500–1000 ms delay, the central
fixation point turned green (go cue (GC)), cueing the subject
to move. At this point, the subject used the joystick to guide a
cursor from the center of the monitor toward the green target.
Once the target was reached, either a juice reward was given (in
the primate case) or a tone sounded indicating that the subject
had successfully completed the task (human case), and the stim-
uli were erased. Subjects were required to return the joystick
to the center position before the next trial started. A schematic
representation of a single trial is shown in Fig. 1.

Table I breaks down the number of trials and neurons
used in our analysis for each PD subject and for our primate
subject.
et al. will be very close to 1 for all times and
\[ P(\lambda) = \lim_{\Delta \to 0} \frac{P(N(t+\Delta) - N(t) = 1|H_t)}{\Delta} \]

where \( H_t \) denotes the history of spikes up to time \( t \). It follows from (1) that the probability of a single spike in a small interval \( (t, t + \Delta] \) is approximately
\[ \Pr(\text{spike in } (t, t + \Delta)|H_t) \cong \lambda(t|H_t)\Delta. \]

Details can be found in [15] and [34]. When \( \Delta \) is small, (2) is approximately the spiking propensity at time \( t \).

E. Point Process Model of STN Dynamics

We formulate a point process model to relate the spiking propensity of each STN neuron to factors associated with movement direction and features of the neuron’s spiking history. We use the model parameters to analyze oscillations, bursting, and directional tuning modulations across the entire trial and make comparisons between two subject groups.

A point process is a series of 0–1 random events that occur in continuous time. For a neural spike train, the 1-s are individual spike times and the 0-s are the times at which no spikes occur. To define a point process model of neural spiking activity, in this analysis, we consider an observation interval \( (0, T] \), and let \( N(t) \) be the number of spikes counted in interval \( (0, t] \) for \( t \in (0, T] \). A point process model of a neural spike train can be completely characterized by its CIF, \( \lambda(t|H_t) \) defined as follows:
\[ \lambda(t|H_t) = \lim_{\Delta \to 0} \frac{P(N(t+\Delta) - N(t) = 1|H_t)}{\Delta} \]

where \( H_t \) denotes the history of spikes up to time \( t \). It follows from (1) that the probability of a single spike in a small interval \( (t, t + \Delta] \) is approximately
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Details can be found in [15] and [34]. When \( \Delta \) is small, (2) is approximately the spiking propensity at time \( t \).

Table I

<table>
<thead>
<tr>
<th>Subject</th>
<th>Total # of Trials Executed</th>
<th># of Neurons Recorded</th>
<th># of Neurons in Analysis</th>
</tr>
</thead>
<tbody>
<tr>
<td>PD-1</td>
<td>159</td>
<td>5</td>
<td>4</td>
</tr>
<tr>
<td>PD-2</td>
<td>38</td>
<td>3</td>
<td>3</td>
</tr>
<tr>
<td>PD-3</td>
<td>141</td>
<td>6</td>
<td>6</td>
</tr>
<tr>
<td>PD-4</td>
<td>24</td>
<td>7</td>
<td>2</td>
</tr>
<tr>
<td>PD-5</td>
<td>44</td>
<td>4</td>
<td>2</td>
</tr>
<tr>
<td>PD-6</td>
<td>189</td>
<td>7</td>
<td>7</td>
</tr>
<tr>
<td>PD-7</td>
<td>100</td>
<td>5</td>
<td>4</td>
</tr>
<tr>
<td>Primate</td>
<td>896</td>
<td>96</td>
<td>35</td>
</tr>
</tbody>
</table>

The CIF generalizes the rate function of a Poisson process to a rate function that is history dependent. Because the conditional intensity function completely characterizes a spike train, defining a model for the CIF defines a model for the spike train [12], [13]. For our analyses, we use the GLM to define our CIF models by expressing for each neuron, the log of its CIF in terms of the neurons spike history and relevant movement covariates [35]. The GLM is an extension of the multiple linear regression model in which the variable being predicted, in this case spike times, need not be Gaussian [27]. GLM also provides an efficient computational scheme for model parameter estimation and a likelihood framework for conducting statistical inferences [12].

We express the CIF for each neuron as a function of movement direction, which corresponds to up, right, left, and down, and the neuron’s spiking history in the preceding 150 ms. Instead of estimating the CIF continuously throughout the entire trial, we estimate it over 350 ms time windows around key epochs and at discrete time intervals each 1 ms in duration. Specifically, we estimate the CIF over 350 ms windows centered at the gray arrow onset (GA), TC onset, GC onset, and movement onset (MV) onsets. Fig. 2 shows all of the time periods for which we estimate the CIF. Henceforth, we omit the superscripts denoting the epoch for a simpler read and express the CIF as follows:
\[ \lambda(t|H_t, \Theta) = \lambda^S(t|\Theta)\lambda^H(t|H_t, \Theta) \]

where \( \lambda^S(t|\Theta) \) describes the effect of the movement direction stimulus on the neural response and \( \lambda^H(t|H_t, \Theta) \) describes the effect of spiking history on the neural response. \( \Theta \) is a parameter vector to be estimated from data. The units of \( \lambda^S(t|\Theta) \) is spikes per second and \( \lambda^H(t|H_t, \Theta) \) is dimensionless. The idea to express the CIF as a product of a stimulus component and a temporal or spike history component was first suggested in [25] and is appealing as it allows one to assess how much each component contributes to the spiking propensity of the neuron. If spiking history is not a factor associated with neural response, then \( \lambda^H(t|H_t, \Theta) \) will be very close to 1 for all times and (3) reduces to an inhomogeneous Poisson process.

The model of the stimulus effect is as follows:
\[ \log \lambda^S(t|\alpha) = \sum_{d=1}^{I} \alpha_d I_d(t) \]

where
\[ I_d(t) = \begin{cases} 1, & \text{for all } t \text{ if movement is in direction } d \\ 0, & \text{otherwise} \end{cases} \]
The \( \{ \alpha_d \}_{d=1}^4 \) parameters measure the effects of movement direction on the spiking propensity. For example, if \( \alpha_1 \) is significantly larger than \( \alpha_2, \alpha_3, \) and \( \alpha_4 \) during movement, then the probability that a neuron will spike is greater when the patient moves in the up direction, suggesting that the neuron may be tuned in the up direction.

Our model of spike history effect is as follows:

\[
\log \lambda_H (t|H_t, \beta, \gamma) = \sum_{j=1}^{10} \beta_j n(t - j : t - (j + 1)) + \sum_{k=1}^{14} \gamma_k n(t - (10k + 9) : t - 10k) \quad (5)
\]

where \( n(a : b) \) is the number of spikes observed in the time interval \([a, b]\) during the epoch. The \( \{ \beta_j \}_{j=1}^{10} \) parameters measure the effects of spiking history in the previous 10 ms, and therefore, can capture refractoriness and/or bursting on the spiking probability in the given epoch. For example, if \( e^{\beta_1} \) is close to zero for any given epoch, then for any given time \( t \), if the neuron had a spike in the previous millisecond then the probability that it will spike again is also close to zero (due to refractory period). Or if \( e^{\beta_5} \) is significantly larger than 1, then for any time \( t \), if the neuron had a spike five prior to \( t \), then the probability that it will spike again is modulated up, suggesting bursting.

The \( \{ \gamma_k \}_{k=1}^{14} \) parameters measure the effects of the spiking history in the previous 10–150 ms on the spiking probability, which may be associated with not only the neuron’s individual spiking activity, but also that of its local neural network. For example, if \( e^{\gamma_1} \) is significantly larger than 1, then for any time \( t \) if the neuron had one or more spikes between 40–50 prior to \( t \), then the probability that it will spike again is modulated up, suggesting 20–25 Hz oscillations.

By combining (4) and (5), we see that the CIF may be written as follows:

\[
\lambda(t|H_t, \Theta) = \exp \left\{ \sum_{d=1}^{4} \alpha_d I_d(t) \right\} \exp \left\{ \sum_{j=1}^{10} \beta_j n(t - (j) : t - (j + 1)) \right\} \exp \left\{ \sum_{k=1}^{14} \gamma_k n(t - (10k + 9) : t10K) \right\} .
\]

(6)

The model parameter vector \( \Theta = \{ \alpha_b, \beta_j, \gamma_k \} \) contains 28 unknown parameters (for each epoch and time window modeled). We compute maximum-likelihood (ML) estimates for \( \Theta \) and 95% confidence intervals of \( \Theta \) for each neuron using glmfit.m in MATLAB.

Finally, to choose the model defined by (6), we varied the history bins \([a \text{ and } b \text{ in } n(a:b)]\) and also varied the basis functions \( I_d(t) \) multiplying the parameters \( \{ \alpha_d \} \) and analyzed the tradeoff between the number of parameters in the model and the ML cost of the fitted model. That is, we computed the Akaike’s criterion (AIC) [3], which is \( 2n \text{ (number of parameters} - \text{log likelihood of model)} \), for each possible model. The model given by (6) was selected as optimal because it rendered the minimum AIC amongst model classes explored.

### F. Model Fitting

Establishing the degree of agreement between a point process model and observations of the spike train and associated experimental variables is a prerequisite for using the point process analysis to make scientific inferences. We used Kolmogorov–Smirnov (KS) plots based on the time-rescaling theorem to assess model goodness-of-fit. The time-rescaling theorem is a well known result in probability theory, which states that any point process with an integrable conditional intensity function may be transformed into a Poisson process with unit rate [24]. A KS plot, which plots the empirical cumulative distribution function of the transformed spike times versus the cumulative distribution function of a unit rate exponential, is used to visualize the goodness-of-fit for each model. The model is better if its corresponding KS plot lies near the 45° line. We computed 95% confidence bounds for the degree of agreement using the distribution of the KS statistic [24]. If a model’s KS plot was within the 95% confidence bounds, we included it in our analyses.

### III. Results

As mentioned earlier, we built point process models for STN neurons in seven Parkinson’s patients and one healthy primate, which captured dynamics across four different epochs within a directed hand-movement task. We summarize results for each species later. For the PD data, 28 STN neuron models passed the KS test and for the primate data, 35 models passed the KS test.

Recall from (2) that \( \lambda(t|H_t) \Delta \) is approximately the probability that the neuron will spike at time \( t \) given extrinsic and intrinsic dynamics up to time \( t \), which is captured in \( H_t \). By virtue of (6), we allow the probability that each STN neuron will spike at some time \( t \) to be modulated by movement direction, short-term history, and long-term history spiking dynamics. Fig. 3 illustrates these three modulation factors on spiking activity for both PD and primate single neuron models by plotting the optimal parameters and their corresponding 95% confidence bounds before and after MV onset. We make the following observations.

1) **Refractoriness:** As illustrated in the second row of Fig. 3, both the PD and primate STN neuron exhibits refractory periods [9] as indicated by down modulation by a factor of ten or more due to a spike occurring 1 ms prior to a given time \( t \). That is, if a spike occurs 1 ms prior to time \( t \), then it is very unlikely that another spike will occur at time \( t \) (\( e^{\beta_1} \leq 1 \) for all \( e^{\beta_1} \) within its 95% confidence band).

2) **Bursting:** As illustrated in the second row of Fig. 3, the PD STN neuron spikes in rapid succession before and after MV onset as indicated by one or more of the short-term history parameters \( e^{\beta_1} \) corresponding to 2–10 ms in the past being larger than 1. That is, if a spike occurs 2–10 ms prior to time \( t \), then it is more likely that another spike will occur at time \( t \). Formally, a neuron bursts if its model parameters satisfy the following: for at least one
Fig. 3. Optimal model parameters for an STN neuron during MV− and MV+ periods of a (left) PD patient and (right) healthy primate. [Top row (movement direction modulation)] Optimal extrinsic factors $e^{\alpha_d}$ for $d = 1, 2, 3, 4(U, R, D, L)$ are plotted in black lines from left to right and corresponding 95% confidence intervals are shaded around each black line in a gray. [Middle row (short-term history modulation)] Optimal short-term history factors $e^{\beta_i}$ for $i = 1, 2, ..., 10$ are plotted in black from right to left and the corresponding 95% confidence intervals are shaded in gray. [Bottom row (long-term history modulation)] Optimal long-term history factors $e^{\gamma_j}$ for $j = 1, 2, ..., 14$ are plotted in black from right to left and corresponding 95% confidence intervals are shaded in gray.

3) 10–30 Hz oscillations: As illustrated in the third row of Fig. 3, the PD STN neuron exhibits 10–30 Hz oscillatory firing before movement. That is, the probability that the PD STN neuron will spike at a given time $t$ is modulated up if a spike occurs 30–100 ms prior to $t$. Formally, a neuron has 10–30 Hz oscillations if its model parameters satisfy the following: for at least one $i = 2, 3, ..., 5$, $LB_i \geq 1$ and $UB_i \geq 1.5$, where $LB_i \leq e^{\beta_i} \leq UB_i$. LB and UB are the 95% lower and upper confidence bounds, respectively.

4) Directional tuning: As illustrated in the first row of Fig. 3, the PD STN neuron appears to exhibit more directional tuning after MV onset. That is, the PD neuron seems more likely to spike in one direction more than at least one other direction. To quantify directional tuning, we performed the following test for each neuron, each time relative to onset, and each epoch:

a) For each direction $d^* = \{U, R, D, L\}$, compute $p_{d^*d} = \Pr(e^{\alpha_{d^*}} > e^{\alpha_d}) = \Pr(\alpha_{d^*} > \alpha_d)$ for $d \neq d^*$. Define $p_{d^*d} = 0$. Use the Gaussian approximation for $\alpha_d$, which is one of the asymptotic properties of ML estimates to compute $p_{d^*d}$.

b) If $\max_{d=1,2,3,4} p_{d^*d} \geq 0.975$ then neuron exhibits directional tuning.

Table II provides a population summary for each of these spiking characteristics for each epoch and subject group.

![Fig. 4. STN population summary using point process model parameters. Dashed line: bursting, dotted line: 10–30 Hz oscillations, and solid line: directional tuning.](image-url)
from one of the four directions, we 

\( p = 0.155 \). 

equals the total number of observa-

tions in the 10–30 Hz range as the integral of the spectrogram in the

oscillation density of the spectrogram

appears), each neuron and each trial. Then, for each spectro-

togram, we computed the oscillation density of the spectrogram

each epoch window (e.g., 350 ms window right before the GA

Tuning increased further in the primate neurons after the TC

is shown in the primate case (see " + " symbols in the

We make the following observations from Fig. 4 and Table II. 

Most neurons in both species exhibit refractoriness. Bursting

is prevalent across all epochs in neural activity of PD patients 

(on average 39% of PD STN neurons burst). In contrast, neural

activity in the healthy primate exhibits little bursting (14% on

average) across all epochs. Ten to thirty hertz oscillations are 

prevalent in neural activity of PD patients during across all

epochs (on average 36%) and significantly decrease relative to 

this baseline post movement as denoted by " + " symbols at the top of Fig. 4. Beta oscillations have been observed experimentally in both Parkinsonian primates and PD patients [7], [8], [10], [17], [29], [32], and attenuation of these oscillation post movement have also been observed [4], [36]. In contrast, an average of 12% of the primate neurons exhibit 10–30 Hz oscillations, which does not significantly modulate

the healthy primate across the trial. Directional tuning is more prevalent in the healthy primate across the trial. In particular, directional tuning increases significantly above baseline right after the GA

is shown in the primate case (see " + " in solid curve at the bottom of Fig. 4 at GA+). This makes sense as the primate knows and moves to one of the four possible directions shown. Tuning increased further in the primate neurons after the TC

appears, as now the subject knows which direction to move when 
cued to move. In contrast, directional tuning fails to increase 
significantly above baseline until right before MV onset (see 

" + " in solid curve at the top plot of Fig. 4 at MV→) in PD STN 
nurons. The lack of significant increase in directional tuning in 

PD STN neurons early on in the trial may reflect the lack of a 

dynamic range in the STN neurons of PD patients, which may 

cause their slow and impaired movements.

For comparison, we also computed spiking characteristics us-

ing traditional methods. Next, we describe these computations.

A. Beta Oscillations

To analyze beta oscillations, we computed spectrograms for 
each epoch window (e.g., 350 ms window right before the GA

appears), each neuron and each trial. Then, for each spectro-

gram, we computed the oscillation density of the spectrogram in the

10–30 Hz range divided by the integral of the spectrogram across 

all frequencies. That is, both were double integrals computed

across specified frequencies in the 10–30 Hz range and across 
al time samples in the epoch. Then, for a given neuron and a

given epoch, we computed the fifth percentile across all trials 
as a lower confidence bound on oscillation density (LBOS). We 
determined the neuron as exhibiting 10–30 Hz oscillations if 

LBOS > 0.155.

B. Bursting

To analyze bursting in these neurons, we computed interspike 
interval (ISI) histograms across each epoch during the trial for 

all neurons and all trials. We then normalized each histogram, 

so that it is summed to 1, and then, computed the bursting density of the histogram in the 2–10 ms range by taking the 

sum of the normalized histogram in the 2–10 ms range. Once 

we computed densities across all trials for a given neuron and 

epoch, we computed the fifth percentile as a lower confidence 

bound (LBBU). For a given neuron and epoch, we determined 

that the neuron bursts if the LBBU > 0.15.

C. Directional Tuning

To analyze directional tuning in these neurons, we computed tuning vectors [20], [21], [33] across each epoch during the 

trial for all neurons and all trials. If the vector sum in all four 
directions lies within 20° from one of the four directions, we 
determined that the neuron is directionally tuned.

The population summary using traditional statistics is shown 

in Fig. 5 for each subject group. When comparing Figs. 4 and 

5, we see similar trends in the spiking characteristics of primate 

STN neurons though absolute percentages slightly differ. In 

particular, we see that for the primate, we have a steady average 
of 17% neurons bursting and 19% neurons oscillating in the 

10–30 Hz range over the entire trial. We also see significantly 

increased directional tuning relative to baseline right after the 

TC is shown (solid curve at the bottom of Fig. 4), whereas, the

---

Fig. 5. STN population summary using traditional statistics. Dashed line: bursting, dotted line: 10–30 Hz oscillations, and solid line: directional tuning.
tuning vector analysis shows a decrease in directional tuning around the GC epoch (solid curve at the bottom of Fig. 5). The reason for this discrepancy is due to the fact that tuning vectors are only capturing first-order statistics of the point process. The point process model parameters take into account the stimulus parameters (\(\alpha's\)) probability distributions (not just the mean values), and directional tuning is determined from these distributions as described earlier. Therefore, making inferences from average tuning vectors can be misleading.

In the human case, we see significant differences between the two analyses. The point process models show significant bursting and oscillations throughout the trial (an average of 39% and 36%, respectively), while traditional methods lead us to believe that there is much less bursting than 10–30 Hz oscillations (an average of 10% and 37%, respectively). The reason the ISI histogram does not show as much bursting in the neurons is precisely because there are also 10–30 Hz oscillations in the spiking activity. Therefore, there are secondary peaks in the ISI histograms between 30–100 ms range. These secondary peaks result in less bursting density, leading us to believe that bursting may not be prevalent. The fact is, PD STN neurons often burst and oscillate in the beta frequency range and this bursting may not be prevalent. The fact is, PD STN neurons is precisely because there are also 10–30 Hz oscillations in the subthalamic nucleus.

Another drawback of using traditional statistics is that they are significantly different from one another, and therefore, using them to define whether a neuron bursts, oscillates, or exhibits directional tuning over a certain epoch is not straightforward. In fact, the population summary shown in Fig. 5 varies significantly as we change thresholds. In contrast, for point process models, we can use the same threshold to determine whether a neuron oscillates and bursts as the threshold is on how model parameters modulate the overall probability that the neuron spikes at any given time.

IV. CONCLUSION

We have applied the point process framework to the analysis of STN microelectrode recordings from PD patients and a healthy nonhuman primate, to understand the relative importance of movement and spiking history on neural responses. We used GLM representations of the point process CIF to develop an efficient likelihood-based approach to model fitting, goodness-of-fit assessment, and inference. The point process model parameters allowed us to identify pathological characteristics of the STN neurons in PD patients, including bursting, 10–30 Hz oscillations, and decreased directional tuning prior to movement. These characteristics, which differ from the characteristics of the non-PD STN neurons, had been previously described using traditional methods. However, such techniques can lead to erroneous inferences when spiking data contains significant temporal dependencies as is the case of PD STN spiking activity. The point process framework is, therefore, a useful paradigm for providing a succinct, quantitative characterization of the pathological behavior of STN spiking activity in PD patients.

**REFERENCES**


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