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Measurement of the time-dependent *CP* asymmetry of partially reconstructed $B^0 \rightarrow D^{*+}D^{*-}$ decays

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National Accelerator Laboratory. Using the technique of partial reconstruction, we measure the timedependent *CP* asymmetry parameters $S = -0.34 \pm 0.12 \pm 0.05$ and $C = +0.15 \pm 0.09 \pm 0.04$. Using the value for the *CP*-odd fraction $R_{\perp} = 0.158 \pm 0.028 \pm 0.006$, previously measured by *BABAR* with fully reconstructed $B^0 \rightarrow D^{*+}D^{*-}$ events, we extract the *CP*-even components $S_+ = -0.49 \pm 0.18 \pm$ 0.07 ± 0.04 and $C_+ = +0.15 \pm 0.09 \pm 0.04$. In each case, the first uncertainty is statistical and the second is systematic; the third uncertainty on S_+ is the contribution from the uncertainty on R_{\perp} . The measured value of the *CP*-even component S_+ is consistent with the value of $\sin 2\beta$ measured in $b \rightarrow$ $(c\bar{c})s$ transitions, and with the Standard Model expectation of small penguin contributions.

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I. INTRODUCTION

In the Standard Model (SM), *CP* violation arises from an irreducible complex phase in the 3×3 quark mixing matrix *V* known as the Cabibbo-Kobayashi-Maskawa (CKM) matrix [1,2]. Unitarity of the CKM matrix requires that the relation $V_{ud}V_{ub}^* + V_{cd}V_{cb}^* + V_{td}V_{tb}^* = 0$, which defines the unitarity triangle, be satisfied. The aim of the *B* Factories is to test the unitarity of the CKM matrix by the precise

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measurement of the angles and sides of the above triangle, whose nonvanishing area indicates violation of *CP* symmetry.

Both the *BABAR* and Belle collaborations have measured the *CP* parameter $\sin 2\beta$, where the angle β is defined as $\beta \equiv \arg[-V_{cd}V_{cb}^*/V_{td}V_{tb}^*]$. The most accurate measurements of $\sin 2\beta$ [3–5] use the $b \rightarrow (c\bar{c})s$ transition, in which B^{0} 's decay to charmonium final states. Measurement of $b \rightarrow c\bar{c}d$ transitions such as $B^0 \rightarrow D^{(*)+}D^{(*)-}$ should yield the same value of $\sin 2\beta$ to the extent that the contributions from penguin processes may be neglected.

The leading and subleading order Feynman diagrams contributing to $B^0 \rightarrow D^{(*)+}D^{(*)-}$ decays are shown in Fig. 1. The effect of neglecting the penguin amplitude has been estimated in models based on factorization and heavy quark symmetry, and the corrections are found to be a few percent [6,7]. Loops involving non-SM particles (for example, charged Higgs or super symmetric particles) could increase the contribution from penguin diagrams and introduce additional phases.

In $\Upsilon(4S) \rightarrow B^0 \overline{B}^0$ events the time-dependent decay rate for $B^0 \rightarrow D^{*+} D^{*-}$ is given by

$$P_{\eta}^{S_{\text{tag}}}(\Delta t) = \frac{e^{-|\Delta t|/\tau_b}}{4\tau_b} \cdot [1 + S_{\text{tag}}S_{\eta}\sin(\Delta m_d\Delta t) + S_{\text{tag}}C\cos(\Delta m_d\Delta t)], \qquad (1)$$

where τ_b is the B^0 lifetime averaged over the two mass eigenstates, Δm_d is the $B^0 \bar{B}^0$ mixing frequency, and Δt is the time interval between the $B^0 \rightarrow D^{*+}D^{*-}$ decay ($B_{\rm rec}$) and the decay of the other B ($B_{\rm tag}$) in the event. The parameter $S_{\rm tag} = \pm 1(-1)$ in Eq. (1) indicates the flavor of the $B_{\rm tag}$ as a B^0 (\bar{B}^0), while $\eta = \pm 1$ indicates the *CP* eigenvalue of the $B^0 \rightarrow D^{*+}D^{*-}$ final state. The parameters *C* and S_η are given by

$$C = \frac{1 - |\lambda|^2}{1 + |\lambda|^2}; \quad S_{\eta} = -\eta \frac{2\Im m(\lambda)}{1 + |\lambda|^2}; \quad \lambda = \frac{q}{p} \frac{\bar{A}}{A}, \tag{2}$$

where $A(\bar{A})$ is the matrix element of the B^0 (\bar{B}^0) decay and p and q are the coefficients appearing in the expression of the physical mass eigenstates B_L , B_H in terms of the flavor eigenstates B, \bar{B} as



FIG. 1. Leading and subleading order Feynman graphs for the $\bar{B}^0 \rightarrow D^{(*)+}D^{(*)-}$ decays.

Since $B^0 \rightarrow D^{*+}D^{*-}$ is the decay of a scalar to two vector mesons, the final state is a mixture of *CP* eigenstates. The *CP*-odd and *CP*-even fractions have been previously measured from the angular analysis of completely reconstructed events [8,9].

A large deviation of the measured parameter S_{η} in Eq. (2) from the value of $\sin 2\beta$ measured in $b \rightarrow (c\bar{c})s$ transitions or a nonzero value of direct *CP* violation [10–12] would be strong evidence of new physics.

Both the *BABAR* [8] and Belle [9] collaborations have studied the *CP* asymmetries of $B^0 \rightarrow D^{*+}D^{*-}$ decays using fully reconstructed events. In this article we report a new measurement based on the technique of partial reconstruction, which allows us to gain a factor of ≈ 5 in the number of selected signal events with respect to the most recent *BABAR* full reconstruction analysis in Ref. [8]. This result is complementary to the latter measurement, because the statistics used are largely independent of each other.

II. THE BABAR DETECTOR AND DATASET

The data sample used in this analysis has been collected with the *BABAR* detector [13] operating at the PEP-II asymmetric-energy *B* Factory located at the SLAC National Accelerator Laboratory. We have analyzed the full *BABAR* data set collected at the Y(4S) mass peak, $\sqrt{s} =$ 10.58 GeV, corresponding to an integrated luminosity of 429.0 fb⁻¹. In addition, we have used 44.8 fb⁻¹ of data taken off-resonance to evaluate the background from events $e^+e^- \rightarrow q\bar{q}$, where *q* represents a *u*, *d*, *s* or *c* quark ("continuum"). To study backgrounds and validate the analysis procedure, we use a GEANT4-based [14] Monte Carlo simulation in which coherent $B\bar{B}$ production is simulated using the package EVTGEN [15].

The asymmetric energies of the PEP-II beams are an ideal environment to study time-dependent *CP* phenomena in the B^0 - \overline{B}^0 system. The boost of the Y(4*S*) in the laboratory frame by $\beta \gamma = 0.56$ increases the separation between the vertices of the two *B* meson daughters, allowing their precise measurement.

The *BABAR* detector is described in detail in Ref. [13]. We give here only a brief description of the main components and their use in this analysis. Tracking is provided by a five-layer silicon vertex detector (SVT) and a drift chamber (DCH). The SVT provides precise position measurements close to the interaction region that are used in vertex reconstruction and low-momentum track reconstruction. The DCH provides excellent momentum measurement of charged particles.

Particle identification of kaons and pions is obtained from ionization losses in the SVT and DCH and from measurements of photons produced in a ring-imaging Cherenkov light detector (the Detector of Internally Reflected Cherenkov light). A CsI(Tl) crystal-based electromagnetic calorimeter (EMC) enables measurement of

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photon energies and electron identification. These systems all operate inside a 1.5 T superconducting solenoid, whose iron flux return is instrumented for muon detection, initially with resistive plate chambers and more recently with limited streamer tubes [16].

III. ANALYSIS METHOD

A. Partial reconstruction

In the partial reconstruction of a $B^0 \rightarrow D^{*+}D^{*-}$ candidate, we reconstruct fully only one of the two $D^{*\pm}$ mesons in the decay chain $D^* \rightarrow D^0 \pi$ [17], by identifying D^0 candidates in one of four final states: $K\pi$, $K\pi\pi^0$, $K\pi\pi\pi$, $K_S^0\pi\pi$. The vertexing algorithm fits the two-step decay tree simultaneously, correctly calculating correlations among all candidates. In the first three D^0 decay modes, assumed to represent Cabibbo-favored decays, charged kaon tracks are selected using particle identification information from the Detector of Internally Reflected Cherenkov light, SVT and DCH. In the last decay mode, K_S^0 candidates are selected by constraining pairs of oppositely charged tracks to a common vertex.

Since the kinetic energy available in the decay $D^* \rightarrow D^0 \pi$ is small, we combine one reconstructed $D^{*\pm}$ with an oppositely charged low-momentum (slow) pion π_s , assumed to originate from the decay of the unreconstructed $D^{*\mp}$, and evaluate the mass $m_{\rm rec}$ of the recoiling D^0 meson by using the momenta of the two particles. For signal events $m_{\rm rec}$ peaks at the nominal D^0 mass [18] with an rms width of about 3 MeV/ c^2 , while for background events no such peak is visible. Thus, $m_{\rm rec}$ is the primary variable to discriminate signal from background. The calculation of $m_{\rm rec}$ proceeds as follows (refer to Fig. 2 for definitions of the various momenta and angles that we use).

The cosine of the angle between the momenta in the Y(4S) center of mass (CM) frame of the *B* and the reconstructed D^* is readily computed as

$$\cos\theta_{BD^*} = \frac{-M_{B^0}^2 + E_{\rm CM}E_{D^*}}{2p_B|\vec{p}_{D^*}|},\tag{3}$$

where all particle masses are set to their nominal values [18], E_{D^*} and \vec{p}_{D^*} are the measured energy and momentum of the reconstructed D^* in the Y(4S) CM frame, $E_{CM}/2$ is the energy of each beam in the CM frame, and $p_B = \sqrt{E_{CM}^2/4 - M_{B^0}^2}$ is the *B* meson CM momentum. Events are required to be in the physical region $|\cos\theta_{BD^*}| < 1$.

Given $\cos\theta_{BD^*}$ and the measured momenta of the D^* and oppositely charged slow pion, the *B* four-momentum can be calculated up to an unknown azimuthal angle ϕ around \vec{p}_{D^*} . For any chosen value of ϕ , conservation laws determine the unreconstructed D^0 four-momentum $q_D(\phi)$, and one can thus compute the corresponding ϕ -dependent invariant mass $m(\phi) \equiv \sqrt{|q_D(\phi)|^2}$. The value of ϕ is not constrained by kinematics and may be chosen arbitrarily, to



FIG. 2. Momenta and angles in the Y(4S) center of mass frame used in partial reconstruction. The orthogonal axes u and v are normal to the momentum \vec{p}_{D^*} of the reconstructed D^* , and u lies in the plane defined by the momenta of the D^* and slow pion, \vec{p}_{D^*} and \vec{p}_s . The angle ϕ is measured in the u - v plane.

the extent that the shape of the resulting $m(\phi)$ distribution may still be described by the type of functions used in our fits. We have chosen the value $\cos\phi = 0.62$, which is the median of the corresponding Monte Carlo distribution for signal events obtained using generated momenta, and define the recoiling D^0 mass $m_{\rm rec} \equiv m(\cos\phi = 0.62)$. We use the same convention to obtain the direction of the unreconstructed D^0 meson.

B. Backgrounds and event selection

Backgrounds to the $B^0 \rightarrow D^{*+}D^{*-}$ process include the following:

- (i) Combinatorial $B\bar{B}$ background, defined as decays other than $B^0 \rightarrow D^{*+}D^{*-}$, for which the $m_{\rm rec}$ distribution is approximately flat.
- (ii) Peaking *BB* background, defined as decays other than B⁰ → D^{*+}D^{*-}, in which the m_{rec} distribution peaks in the signal region. It will be shown later that the contribution from this background is negligible.
 (iii) Background from non-bb events.

Combinatorial $B\bar{B}$ background events are reduced by the following requirements. For the $K_S^0 \pi \pi$ mode, we require the invariant mass of the pion pair to be within 25 MeV/ c^2 of the K_S^0 mass [18]. The corresponding vertex must be separated by more than 3 mm from the beam axis. For the $K\pi\pi^0$ mode, π^0 candidates are formed from pairs of photons detected in the EMC, with energies greater than 40 MeV, for which the invariant mass differs by less than 20 MeV/ c^2 from the nominal π^0 mass [18]. The reconstructed D^0 mass must be

equal to the nominal one [18] within 2 or 2.5 standard deviations, depending on the D^0 reconstruction mode. The momenta in the Y(4S) CM frame of the reconstructed D^* and π_s from the missing D^0 must be, respectively, in the range 1.3–2.1 GeV/c and smaller than 0.6 GeV/c. The difference $\Delta M = |M_{D^*} - M_{D^0} - M_{\pi}|$ must be equal to the nominal [18] value within 1 or 1.5 MeV/c², according to the presence or absence of DCH hits in the pion track appearing in the reconstructed decay $D^* \rightarrow D^0 \pi$. The probability of the vertex fits must be greater than 10^{-2} , for both the D^0 and the D^* reconstruction.

The requirement on the D^0 vertex fit probability introduces a small but measurable bias toward lower values of the *B* lifetime. Because of the partial reconstruction, the tracks used to make the D^0 vertex may originate from the same or different *B* mesons. In the latter case, since not all tracks are from the same point in space, the χ^2 of the vertex fit tends to be bigger. This effect worsens with increasing distance between the two *B* decay vertices, causing vertices further apart to be rejected more frequently. We have verified this on signal Monte Carlo events, for which we have measured a lifetime lower than the generated value. Consequently, for the signal Δt probability distribution functions (PDF's) we use the value of τ_b fitted to signal Monte Carlo.

In events passing this selection we find more than one candidate decay chain in about 25% of the cases, usually differing only in the slow pion π_s , but sometimes in the components of the reconstructed D^* . When this happens, we choose one candidate chain, based respectively on the largest number of DCH hits in the π_s , or according to a χ^2 based on the reconstructed D^0 mass and ΔM quantity above. For signal Monte Carlo, the probability for this candidate chain to be the correct one is 0.95.

The main suppression of continuum background is obtained by requiring that the ratio R_2 of the 2nd to the 0th Fox-Wolfram moment [19], computed using all charged particles and EMC clusters not matched to tracks, be less than 0.3.

C. Fisher discriminant

To further reduce continuum background, we combine several event-shape variables into a Fisher discriminant [20] F. Discriminating power originates from the observation that $q\bar{q}$ events tend to be jet-like, whereas $B\bar{B}$ events have a more spherical energy distribution. Rather than applying requirements on F, we use the corresponding distribution in the fits described in Sec. III E.

Our Fisher discriminant is a linear combination of variables chosen, according to Monte Carlo studies, to maximize the separation between $B\bar{B}$ and continuum events. The first nine variables describe the energy flow inside nine concentric cones centered around the direction of the reconstructed $D^{*\pm}$. In addition, we use the momenta of the charged and the neutral particle closest to the cone axis,

the polar angles in the CM of the reconstructed D^* momentum and the thrust axis T for charged tracks in the B_{tag} vertex (see next paragraph), the angle between the reconstructed D^* momentum and T, and the sum $S = \sum_i p_i \times P_2(\cos\theta_i)$ over the B_{tag} charged tracks, in which p_i is momentum, P_2 is the 2nd Legendre polynomial of argument $\cos\theta_i$, and θ_i is the angle between track *i* at the origin and T.

D. Flavor tagging and decay time measurement

For this analysis, two measurements are needed: the difference Δt between the proper decay times of the partially reconstructed *B* meson and the other *B* meson in the event, and the flavor of the latter.

The flavor tagging algorithm is based on tracks identified as electrons, muons or kaons. The electron and muon tags contribute equally to the total sample and, since these events are kinematically almost indistinguishable and have very similar effective tagging efficiency, we treat them as one homogeneous "lepton" sample.

The tagging tracks must be chosen among those not used in $B_{\rm rec}$ reconstruction and must originate from within 4 mm (3 cm) of the interaction point in the transverse (longitudinal) view. The momentum of the lepton candidates is required to be greater than 1.1 GeV/c in order to reject most leptons from charmed meson decays. If one or more lepton candidates are qualified, the tag flavor is assigned based on the charge of the lepton with the highest center-of-mass momentum. If two or more qualified kaons are present, the event is used only if the flavor is unambiguous. If both a lepton and a kaon tag are available, the lepton tag is used.

The time difference Δt is calculated using $\Delta t = \Delta z / \gamma \beta c$, where $\Delta z = z_{rec} - z_{tag}$ is the difference between the z-coordinates of the partially reconstructed B_{rec} and B_{tag} vertices and the boost parameters are calculated using the measured beam energies. The uncertainty $\sigma_{\Delta t}$ on Δt is calculated from the results of the z_{rec} and z_{tag} vertex fits. We require $|\Delta t| < 20$ ps and $\sigma_{\Delta t} < 2.5$ ps.

We define the $B_{\rm rec}$ vertex as the decay point of the fully reconstructed $D^{*\pm}$. The π_s track from the other $D^{*\pm}$ is not used, since it undergoes significant multiple Coulomb scattering and hence does not improve the $z_{\rm rec}$ measurement resolution.

The B_{tag} vertex reconstruction depends on the tagging category. For kaon-tagged events, we obtain z_{tag} from a beam spot constrained vertex fit of all charged tracks in the event, excluding those from the B_{rec} meson, and excluding also tracks within 1 rad of the unreconstructed D^0 momentum in the CM frame, which presumably originate from the D^0 decay. We require the probability of this fit to be greater than 10^{-2} . For lepton tagged events, we use the lepton track parameters and errors, and the measured beam spot position and size in the plane perpendicular to the beams (the x-y plane). We find the position of the point in space for which the sum of the χ^2 contributions from the lepton track and the beam spot is minimum. The *z*-coordinate of this point is taken as z_{tag} .

The beam spot is measured on a run-by-run basis using 2-prong events (Bhabha and $\mu^+\mu^-$), and has an rms size of approximately 120 μ m in the horizontal dimension (*x*), 5 μ m in the vertical dimension (*y*), and 8.5 mm along the beam direction (*z*). The average *B* meson flight distance in the *x*-*y* plane is 30 μ m. To account for the *B* flight distance in the beam spot constrained vertex fit, 30 μ m are added in quadrature to the effective *x* and *y* sizes.

E. Probability distribution functions

We use two PDF's, P_{on} for on-resonance, and P_{off} for off-resonance data. The former depends on the variables m_{rec} , F, Δt , $\sigma_{\Delta t}$, S_{tag} , and is given by the sum of the PDF's for the different event types described above,

$$P_{\rm on} = f_{B\bar{B}} [f_{\rm sig} P_{\rm sig} + (1 - f_{\rm sig}) P_{\rm comb}] + (1 - f_{B\bar{B}}) P_{q\bar{q}}, \quad (4)$$

where P_{sig} , P_{comb} , and $P_{q\bar{q}}$ are respectively the PDF's for signal events, for combinatorial background from $B\bar{B}$, and for continuum. Moreover, $f_{B\bar{B}}$ is the fraction of $B\bar{B}$ events in our sample, and f_{sig} is the fraction of signal events in $B\bar{B}$ events. The PDF for off-resonance data, P_{off} , is reduced to just one component, $P_{q\bar{q}}$, as the off-peak sample contains only continuum events.

According to Monte Carlo, the distributions of $B^0 \bar{B}^0$ and $B^+ B^-$ combinatorial background events are very similar and can be described well by the same PDF.

We do not consider the fraction of $B\bar{B}$ events a free parameter, but fix it to $f_{B\bar{B}} = 1 - f_{q\bar{q}}$, where $f_{q\bar{q}}$ is the fraction of continuum events in the on-peak sample and is defined by

$$f_{q\bar{q}} = \frac{N_{\text{off-peak}}}{N_{\text{on-peak}}} \frac{\mathcal{L}_{\text{on-peak}}}{\mathcal{L}_{\text{off-peak}}},$$
(5)

where N's are the number of events left by our selection in the on- and off-peak samples and \mathcal{L} 's are the integrated on- and off-peak luminosities.

Each of the P_i (i = sig, comb, $q\bar{q}$) can be expressed as the product of three one-dimensional PDF's,

$$P_i(m_{\rm rec}, F, \Delta t, \sigma_{\Delta t}, S_{\rm tag}) = \mathcal{M}_i(m_{\rm rec})\mathcal{F}_i(F)T'_i(\Delta t, \sigma_{\Delta t}, S_{\rm tag}),$$
(6)

that are the probability distributions of the recoil D^0 mass $\mathcal{M}_i(m_{\rm rec})$, the Fisher discriminant function $\mathcal{F}_i(F)$, and the decay time difference function $T'_i(\Delta t, \sigma_{\Delta t}, S_{\rm tag})$. This follows from extensive Monte Carlo studies showing that the correlations among these variables are negligible.

1. $\mathcal{M}(m_{rec})$ and $\mathcal{F}(F)$ PDF's

The $m_{\rm rec}$ distribution of all sample components can be well modeled in the lower region of the spectrum with a so called "Argus function" [21],

$$\mathcal{A}(m_{\rm rec}) = m_{\rm rec} \sqrt{1 - (m_{\rm rec}/m_{\rm ep})^2} \cdot e^{c \cdot m_{\rm rec}/m_{\rm ep}}, \qquad (7)$$

where $m_{\rm ep}$ is the kinematic endpoint ($m_{\rm rec} \leq m_{\rm ep}$) and *c* is a free parameter describing the slope. This function alone, however, is not sufficient to account for the abrupt fall of the $m_{\rm rec}$ spectrum near the kinematic endpoint. For the signal sample we model this shoulder with an asymmetric error function with widths σ_l and σ_r , tapered off at low $m_{\rm rec}$ by an exponential factor with decay constant *a*,

$$\begin{aligned} \mathcal{E}(m_{\rm rec}) \\ &= \begin{cases} e^{m_{\rm rec}/a} [1 - \mathrm{erf}(m_{\rm rec} - m_{\rm ep})/(\sqrt{2}\sigma_l)], & m_{\rm rec} < m_{\rm ep} \\ e^{m_{\rm rec}/a} [1 - \mathrm{erf}(m_{\rm rec} - m_{\rm ep})/(\sqrt{2}\sigma_r)], & m_{\rm rec} > m_{\rm ep} \end{cases} \end{aligned}$$

Thus, we describe the signal $m_{\rm rec}$ distribution with a combination of three functions: a Gaussian *G* having average m_G and standard deviation σ_G for the well reconstructed peaking component; an Argus function, mainly for events in which the right D^* is combined with a low-momentum pion from another decay chain; and the \mathcal{E} function

$$\mathcal{M}_{\rm sig}(m_{\rm rec}) = f_{\rm sig}^{\mathcal{A}} \cdot \mathcal{A}(m_{\rm rec}) + (1 - f_{\rm sig}^{\mathcal{A}}) \cdot [f^G \cdot G(m_{\rm rec}) + (1 - f^G) \cdot \mathcal{E}(m_{\rm rec})].$$
(8)

In Eq. (8) $f_{\text{sig}}^{\mathcal{A}}$ is the fraction of events described by the Argus component and f^{G} is the fraction of events in the Gaussian peak relative to the non-Argus component.

For the background, both combinatorial and continuum, we set the fraction of the Gaussian component to zero, and model the distribution at the endpoint with a simple error function of width σ . However, for the case of combinatorial background in kaon-tagged events, we find that two different Argus components (\mathcal{A}_1 and \mathcal{A}_2) are needed to correctly describe the entire reconstructed mass spectrum. We thus define two PDF's according to

$$\mathcal{M}_{\text{comb}}(m_{\text{rec}}) = f_{\text{comb}}^{\text{erf}} \cdot \operatorname{erf}(m_{\text{rec}} - m_{\text{ep}}; \sigma_{\text{comb}}) + (1 - f_{\text{comb}}^{\text{erf}}) \cdot [f_{\text{comb}}^{\mathcal{A}_1} \cdot \mathcal{A}_1(m_{\text{rec}}) + (1 - f_{\text{comb}}^{\mathcal{A}_1}) \cdot \mathcal{A}_2(m_{\text{rec}})], \qquad (9)$$

$$\mathcal{M}_{q\bar{q}}(m_{\rm rec}) = f_{q\bar{q}}^{\rm erf} \cdot \operatorname{erf}(m_{\rm rec} - m_{\rm ep}); \sigma_{q\bar{q}}) + (1 - f_{q\bar{q}}^{\rm erf}) \cdot \mathcal{A}_1(m_{\rm rec}).$$
(10)

The parameter m_{ep} represents simultaneously the two Argus endpoints and the error function inflection point.

The Fisher discriminant PDF \mathcal{F}_i is parametrized by two Gaussian functions for each event type $i = (B\bar{B}, q\bar{q})$, having standard deviations σ_i^L and σ_i^R , and common mean μ_i ,

$$\mathcal{F}_{i}(F) \propto \begin{cases} \exp[-(F - \mu_{i})^{2}/2(\sigma_{i}^{L})^{2}] & F < \mu_{i} \\ \exp[-(F - \mu_{i})^{2}/2(\sigma_{i}^{R})^{2}] & F > \mu_{i}. \end{cases}$$
(11)

Since the Fisher variable is designed to discriminate between $q\bar{q}$ and $B\bar{B}$ events, we expect the Fisher

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discriminant for signal events to be indistinguishable from that of $B\bar{B}$ combinatorial events. We have verified this expectation with Monte Carlo studies, and thus use the same Fisher discriminant to describe both event types.

2. $\Delta t PDF$'s

The Δt -dependent part of the PDF is a convolution of the form

$$T_{i}^{\prime}(\Delta t, \sigma_{\Delta t}, S_{\text{tag}}) = \int d\Delta t_{\text{true}} T_{i}(\Delta t_{\text{true}}, S_{\text{tag}}) \mathcal{R}_{i}(\Delta t - \Delta t_{\text{true}}, \sigma_{\Delta t}), \qquad (12)$$

where *T* is the distribution of Δt_{true} , the true decay time difference, and \mathcal{R} is a resolution function that parametrizes detector resolution and systematic offsets in the measured positions of vertices.

Taking into account the mistag probability and the effect of tags due to the unreconstructed D^0 , the Δt_{true} signal PDF in Eq. (12) can be written as

$$T_{\rm sig} = \frac{1}{4\tau_b} e^{-|\Delta t_{\rm true}|/\tau_b} \cdot \{1 - S_{\rm tag} \Delta \omega (1 - \alpha) + S_{\rm tag} (1 - 2\omega)(1 - \alpha) \cdot [C \cos(\Delta m_d \Delta t_{\rm true}) + S \sin(\Delta m_d \Delta t_{\rm true})]\},$$
(13)

where the time-dependent *CP* asymmetry parameters *S* and *C* are the object of the measurement discussed in the present article and α (see Sec. IV B) is the fraction of events in which the tagging track is from the unreconstructed D^0 . We parametrize possible detector effects leading to a small difference between the mistag probability of B^0 tags (ω^+) and that of \bar{B}^0 tags (ω^-), by using the average mistag rate $\omega \equiv (\omega^+ + \omega^-)/2$ and the mistag rate difference $\Delta \omega \equiv \omega^+ - \omega^-$ as parameters of the PDF.

Since the $B\bar{B}$ combinatorial background is dominated by non-*CP* final states, the *CP* asymmetry is expected to be negligible. However, we allow the PDF to accommodate some contamination from *CP* final states. Therefore, we parametrize the $B\bar{B}$ background Δt_{true} distribution with a PDF similar to that for signal events given in Eq. (13). We also add a fraction f_{δ} of a δ -function, to allow for a zero-lifetime component,

$$T_{\text{comb}} = f_{\text{comb}}^{\delta} \cdot \delta(|\Delta t_{\text{true}}|)(1 - S_{\text{tag}}\Delta\omega_{\text{comb}}^{\delta}) + (1 - f_{\text{comb}}^{\delta}) \cdot \frac{1}{4\tau_{\text{comb}}} e^{-|\Delta t_{\text{true}}|/\tau_{\text{comb}}} \cdot \{1 - S_{\text{tag}}\Delta\omega_{\text{comb}} + S_{\text{tag}} \cdot [C_{\text{comb}}\cos(\Delta m_d\Delta t_{\text{true}}) + S_{\text{comb}}\sin(\Delta m_d\Delta t_{\text{true}})]\}.$$
(14)

The second term of the PDF is obtained from Eq. (13) with $\omega = \alpha = 0$, as these are not defined for background events. The C_{comb} , S_{comb} parameters describe small fluctuations in the Δt_{true} distribution of background events and

possible *CP* event contamination, leading to a small effective *CP* violation value.

The $\Delta \omega$ parameters, which for signal events is the difference in the mistag probabilities for B^0 and \bar{B}^0 , allow for differences in the number of events tagged as a B^0 or \bar{B}^0 in the same background sample. We use this PDF to describe both the $B^0\bar{B}^0$ and B^+B^- components.

The PDF for the background due to continuum events is modeled with a simple exponential decay distribution plus a fraction f_{δ} of a δ -function,

$$T_{q\bar{q}} = f_{q\bar{q}}^{\delta} \cdot (1 - S_{\text{tag}} \Delta \omega_{q\bar{q}}^{\delta}) \cdot \delta(|\Delta t_{\text{true}}|) + (1 - f_{q\bar{q}}^{\delta}) \cdot (1 - S_{\text{tag}} \Delta \omega_{q\bar{q}}) \cdot \frac{1}{4\tau_{q\bar{q}}} e^{-|\Delta t_{\text{true}}|/\tau_{q\bar{q}}},$$
(15)

where the parameters $\Delta \omega_{q\bar{q}}^{\delta}$ and $\Delta \omega_{q\bar{q}}$ allow for differences in the number of events tagged as a B^0 or \bar{B}^0 in this sample.

3. Resolution functions

The functions T'_i of the measured time difference Δt , to be used in the fits, are obtained by convolving the T_i PDF's of Eqs. (13)–(15), with the appropriate resolution function for events of type i (i = sig, comb, $q\bar{q}$).

The resolution functions are parametrized as the sum of three Gaussian functions,

$$\mathcal{R}_{i}(t_{r}, \sigma_{\Delta t}) = f_{i}^{n} \mathcal{G}_{i}^{n}(t_{r}, \sigma_{\Delta t}) + (1 - f_{i}^{n} - f_{i}^{o}) \mathcal{G}_{i}^{w}(t_{r}, \sigma_{\Delta t}) + f_{i}^{o} \mathcal{G}_{i}^{o}(t_{r}),$$
(16)

where $t_r = \Delta t - \Delta t_{true}$ is the residual of the Δt measurement, and \mathcal{G}_i^n , \mathcal{G}_i^w , and \mathcal{G}_i^o are the "narrow," "wide," and "outlier" Gaussian functions. The narrow and wide Gaussian functions incorporate information from the Δt uncertainty $\sigma_{\Delta t}$, and account for systematic offsets in the estimation of $\sigma_{\Delta t}$ and the Δt measurement. They have the form

$$\mathcal{G}_{i}^{k}(t_{r},\sigma_{\Delta t}) \equiv \frac{1}{\sqrt{2\pi}s_{i}^{k}\sigma_{\Delta t}} \cdot \exp\left(-\frac{(t_{r}-b_{i}^{k}\sigma_{\Delta t})^{2}}{2(s_{i}^{k}\sigma_{\Delta t})^{2}}\right), \quad (17)$$

where the index k takes the values k = n, w for the narrow and wide Gaussian functions, and b_i^k and s_i^k are parameters determined by fits. The outlier Gaussian function, describing a small fraction of events with badly measured Δt , has the form

$$\mathcal{G}_{i}^{o}(t_{r}) \equiv \frac{1}{\sqrt{2\pi}s_{i}^{o}} \exp\left(-\frac{(t_{r}-b_{i}^{o})^{2}}{2(s_{i}^{o})^{2}}\right).$$
 (18)

In all fits, the values of b_i^o and s_i^o are fixed to 0 and 8 ps, respectively, and are later varied to evaluate systematic uncertainties.

F. Analysis procedure

After the event selection described in Sec. III B is complete, the rest of the analysis proceeds with a series of unbinned maximum-likelihood fits, performed simultaneously on the on- and off-resonance data samples and independently for the lepton tagged and kaon-tagged events. The procedure can be logically divided in the following three steps, which we shall discuss in detail in the following paragraphs:

(1) In the first step we determine the signal fraction f_{sig} in Eq. (4) and the shape of $\mathcal{M}(m_{rec})$ and $\mathcal{F}(F)$ in Eq. (6) for the different classes of events (signal and backgrounds, kaon and lepton tagging categories). This is done by fitting data with the PDF

$$P_i(m_{\rm rec}, F) = \mathcal{M}_i(m_{\rm rec})\mathcal{F}_i(F), \qquad (19)$$

ignoring the time dependence; we refer to this step as the kinematic fit.

- (2) In the second step we determine the tagging dilution due to wrong tag assignments.
- (3) In the last step we perform the time-dependent fit to the data. We fix all parameter values obtained in the previous steps and use the full PDF of Eq. (6) to determine the parameters of the resolution functions, $T'_i(\Delta t, \sigma_{\Delta t}, S_{tag})$, and the *CP* asymmetry values *C*, *S* of the signal and of the *B* \bar{B} combinatorial background component.

The fitting procedure has been validated using both full Monte Carlo and, where the requested number of events would be too large, the technique of "toy" Monte Carlo. In a toy Monte Carlo, events are described by a small number of variables which are generated according to our PDF's.

IV. RESULTS

Event selection yields the numbers of events listed in the top two rows of Table I. The third and fourth rows show the number of continuum and $B\bar{B}$ events calculated, using Eq. (5), from the number of off-peak events in the second row. The numbers of signal events in the last line of the

TABLE I. Event selection yield. The first uncertainty shown is statistical, while the second uncertainty on the number of continuum events accounts for a 1% relative uncertainty on the on-peak and off-peak luminosities.

	Number of events		
	Kaon tag	Lepton tag	
On-peak	61179	20855	
Off-peak	1025	51	
Continuum	$9814 \pm 307 \pm 196$	$488\pm68\pm10$	
$B\bar{B}$	51365 ± 364	20367 ± 69	
N _{sig}	3843 ± 397	1129 ± 218	

table are calculated using the signal fractions obtained from the kinematic fit described in the next section.

A. Kinematic fit

We begin by fitting the shape of our signal, $\mathcal{M}_{sig}(m_{rec})$, using a large sample of Monte Carlo signal events. The parameters most relevant to determine directly the signal fraction in the data, and consequently our final result for *S* and *C*, will be released again in the final kinematic fit. They are [refer to Eq. (8)]: the Gaussian fraction f_G , mean value m_G , and standard deviation σ_G , and are shown in the last section of Table II.

Next we fit the Fisher $\mathcal{F}_{q\bar{q}}$ and recoil mass $\mathcal{M}_{q\bar{q}}$ distribution to the off-peak data sample. As the number of offresonance events selected in the lepton tagged sample is too small to yield convergence, we set the lepton tag sample parameters to the corresponding values obtained from the fit to the kaon tag sample. Because of the small continuum fraction in the lepton sample, we judge that this does not introduce any significant systematic effect. The $\mathcal{F}_{a\bar{a}}$ parameters are fixed in all subsequent fits.

We initialize the parameters of the $B\bar{B}$ combinatorial background PDF directly from the data, using a sample of events in which the contribution of signal events is much reduced. We obtain this sample by combining a D^* with a pion of wrong sign charge (WS sample). We have verified, both on Monte Carlo and in the $m_{\rm rec}$ sideband for data (1.836–1.856 GeV/ c^2), that the shape of the $\mathcal{M}(m_{\rm rec})$ distribution for combinatorial $B\bar{B}$ background is well described by that of the WS data sample.

To evaluate a possible contribution from a peaking component in the $B\bar{B}$ background events, we have allowed the Gaussian fraction f_G in Eq. (8) to float in a fit to a sample of Monte Carlo background events; this fraction is found to be 0.000 ± 0.002 , and is therefore set to zero.

Finally we fit the on-peak data sample, leaving as free parameters the fraction f_{sig} of signal events in the $B\bar{B}$ component, some of the shape parameters of the continuum and $B\bar{B}$ combinatorial background \mathcal{M}_{comb} , some of the signal parameters in \mathcal{M}_{sig} , and the shape parameters of the Fisher discriminant $\mathcal{F}_{B\bar{B}}$. Table II summarizes the results and provides information about which parameters are released in the fit (statistical uncertainties given) and which ones are taken from previous fits (no uncertainty given).

The final results of the kinematic fits for the kaon and lepton tagged sample are shown in Figs. 3 and 4.

B. Determination of mistag probabilities

A common problem of analyses using the partial reconstruction technique is that a fraction of the tracks used in tagging may belong to the unreconstructed D^0 , leading to a mistag of the event. As the tracks originating from the missing D^0 tend to align to its direction of flight, this fraction can be reduced by applying a constraint on the

PDF	Parameter	Description	Kaon tags	Lepton tags
$\mathcal{F}_{B\bar{B}}(F)$	$f_{\rm sig}$	Signal fraction	$(7.5 \pm 0.7)\%$	$(5.5 \pm 1.1)\%$
	$\mu_{Bar{B}}$		0.723 ± 0.005	0.721 ± 0.009
	$\sigma^L_{_{Rar{B}}}$		0.361 ± 0.003	0.380 ± 0.006
	$\sigma^{R}_{_{R\bar{R}}}$		0.469 ± 0.004	0.532 ± 0.006
$\mathcal{M}_{a\bar{a}}(m_{\rm rec})$	f^{A}	Argus fraction	1.0	1.0
11	m _{ep}	Argus endpoint (GeV/ c^2)	1.8696	1.8696
	c	Argus exponent	-17 ± 9	-17
	$f_{\rm comb}$	erf fraction	0.52 ± 0.15	0.25
	$\sigma_{ m comb}$	erf width (GeV/ c^2)	0.0005 ± 0.0002	0.0005
$\mathcal{M}_{\rm comb}(m_{\rm rec})$	$f^{\mathcal{A}_1}$	Argus fraction	0.27 ± 0.06	1.0
	m _{ep}	Argus end point (GeV/ c^2)	1.8696	1.8695
	c_1	\mathcal{A}_1 exponent	-49 ± 38	-15 ± 2
	c_2	\mathcal{A}_2 exponent	-0.56 ± 0.25	• • •
	$f_{a\bar{a}}$	erf fraction	0.26 ± 0.04	0.41 ± 0.06
	$\sigma_{a\bar{a}}$	erf width (GeV/ c^2)	$(75 \pm 9) \cdot 10^{-5}$	$(72 \pm 2) \cdot 10^{-5}$
$\mathcal{M}_{\rm sig}(m_{\rm rec})$	f_{G}^{II}	Gaussian fraction	0.46 ± 0.06	0.64 ± 0.12
6	m_G	Gaussian peak (GeV/ c^2)	1.8638 ± 0.0002	1.8635 ± 0.0003
	σ_G	Gaussian width (GeV/ c^2)	0.0017 ± 0.0002	0.0019 ± 0.0003

TABLE II. Results of the final kinematic fits. The values of fixed parameters are given without uncertainties.

cosine of the CM opening angle θ_{tag} between the tagging track and the direction of the unreconstructed D^0 . We require $\cos\theta_{\text{tag}} \leq 0.75(0.50)$ for the kaon- (lepton) tagged sample.

We find from signal Monte Carlo that before this requirement 26% (13%) of kaon (lepton) tags originate

from a missing D^0 . We call these events "*D*-tags," while "*T*-tags" are those in which the tagging track (either direct or cascade) is from the tag *B*.

To reduce dependency on Monte Carlo the fraction α of *D*-tags remaining after the $\cos\theta_{\text{tag}}$ constraint is measured using data, as explained below.

Defining the total number N_T (N_D) of T (D) tags, the number N_T^l (N_D^l) of them that satisfy the $\cos\theta_{\text{tag}}$ requirement, and the number N_T^g (N_D^g) of them that do not, α is given by



FIG. 3 (color online). Result of the kinematical fit of kaon (top) and lepton tagged (bottom) data events, with PDF's overlaid: total PDF (solid line), total background ($B\bar{B}$ + continuum, short dashed line), $B\bar{B}$ combinatorial background (dashed line), continuum *u*, *d*, *s*, *c* background (dot-dashed line) and signal (red, dotted line).



FIG. 4. Result of the kinematical fit of kaon- (top) and lepton tagged (bottom) events for the Fisher function, with PDF's overlaid: total $B\bar{B}$ (solid line), $B\bar{B}$ (dashed line) and continuum u, d, s, c background (dash-dotted line).

TABLE III. Values of mistag parameters and α used in the final fit. The *b* lifetime values were obtained from the fit of signal Monte Carlo. The statistical uncertainties are given.

Parameter	Kaon tags	Lepton tags
ω	0.201 ± 0.002	0.104 ± 0.002
$\Delta \omega$	-0.011 ± 0.003	0.001 ± 0.005
α	0.12 ± 0.04	0.0 ± 0.02
τ_b (ps)	1.458 ± 0.014	1.518 ± 0.018
$\Delta m_d ~(\mathrm{ps}^{-1})$	0.507 =	± 0.004

$$\alpha = N_D^l / (N_D^l + N_T^l) = \frac{f^D p^D}{(f^D p^D + (1 - f^D) p^T)},$$
 (20)

where $p^T = N_T^l/N_T$ $(p^D = N_D^l/N_D)$ is the probability, taken from signal Monte Carlo, for a *T*-tag (*D*-tag) to be from a track satisfying the $\cos\theta_{\text{tag}}$ cut. The fraction of *D*-tags, $f^D = N_D/(N_T + N_D)$, is given by

$$f^D = (p^T - f^l)/(p^T - p^D).$$
 (21)

The fraction $f^l = N^l/(N^l + N^g)$ is obtained from the kinematic fit of the data: N^l is the number of signal events that have $\cos\theta_{\text{tag}} \le 0.75(0.50)$ and N^g is the number of signal events with $\cos\theta_{\text{tag}} \ge 0.75(0.50)$ for kaon (lepton) tag events.

In this way we obtain the values $\alpha = 0.12 \pm 0.04$ for kaon tags and $\alpha = 0.00 \pm 0.02$ for lepton tags, as shown in Table III, where we also list the mistag parameters ω and $\Delta \omega$, α , τ_b , and Δm_d that we will need in the final Δt fit. We use information from signal Monte Carlo events to determine the mistag probability $\omega = 0.201 \pm 0.002(0.101 \pm 0.002)$ and mistag difference $\Delta \omega = -0.011 \pm 0.003(0.001 \pm 0.005)$ for the kaon- (lepton) tagged samples. We use the world average value for Δm_d [18], and the values fitted to signal Monte Carlo for τ_b , as discussed in Sec. III B.

C. Time dependent fit

After the kinematic fit is complete and mistag probabilities are determined, we can proceed to the final Δt fit to extract the *CP*-violating parameters *S* and *C*.

We start by fitting the signal Δt PDF and its resolution function using a pure signal Monte Carlo sample; the parameters determined in this way will be used to initialize the signal PDF, and some of them will be left free again in the final Δt fit.

Next we fit the resolution function parameters and the effective lifetime of the continuum background, using the off-peak data sample. For the kaon tag sample, the data strongly disfavor a component with nonzero lifetime, therefore we fix $f_{\delta} = 1$, and leave free in the final Δt fit only $\Delta \omega_{\delta}$ from Eq. (15). For lepton tags, as the real data sample is too small, we obtain resolution and Δt parameters from continuum Monte Carlo.

We use the continuum parameters obtained above in the next fit stage, where we obtain the $B\bar{B}$ background resolution function and Δt parameters using real data, by restricting the fit to events in a sideband region (1.836–1.856 GeV/ c^2) of

TABLE IV. Result of the final full fit. The values of fixed parameters are given without uncertainties.

PDF	Parameter	Description	Kaon tags	Lepton tags
$B\bar{B}$ resolution model	b^n	Offset of narrow Gaussian	-0.16 ± 0.01	-0.022 ± 0.014
	b^o	Offset of outlier Gaussian (ps)	0.0	0.0
	b^w	Offset of wide Gaussian	-1.0 ± 0.2	-0.7 ± 0.7
	f^n	Fraction of narrow Gaussian	0.93 ± 0.01	0.977 ± 0.004
	f^o	Fraction of outlier Gaussian	0.008 ± 0.001	0.006 ± 0.002
	s^n	See Eq. (17)	1.03 ± 0.03	1.02 ± 0.02
	s ^o	See Eq. (18) (ps)	8.0	8.0
	s^w	See Eq. (17)	3.0	5.6
Continuum Δt	$\Delta \omega^{\delta}_{aar{a}}$	See Eq. (15)	-0.04 ± 0.02	0.03
$B\bar{B} \ \Delta t$	$f_{\rm comb}^{\delta}$	See Eq. (14)	0.10 ± 0.02	0.25 ± 0.02
	$\Delta \omega_{\rm comb}^{\delta}$	See Eq. (14)	0.04 ± 0.12	-0.08 ± 0.07
	$\Delta \omega_{\rm comb}$	See Eq. (14)	-0.025 ± 0.012	0.012 ± 0.021
	$ au_{ m comb}$	Effective lifetime (ps)	1.318 ± 0.023	1.272 ± 0.031
	$C_{\rm comb}$	Cosine coefficient	-0.022 ± 0.024	-0.024 ± 0.041
	$S_{\rm comb}$	Sine coefficient	0.004 ± 0.014	-0.023 ± 0.024
Signal resolution model	b^n	Offset of narrow Gaussian	-0.35 ± 0.09	-0.3 ± 0.2
	b^w	Offset of wide Gaussian	8 ± 3	• • •
	f^n	Fraction of narrow Gaussian	0.992 ± 0.007	1.0 ± 0.1
	f^o	Fraction of outlier Gaussian	0.0	• • •
	S^n	See Eq. (17)	1.13 ± 0.12	1.17 ± 0.21
	s^w	See Eq. (17)	2.6	• • •
Signal Δt	С	_	$+0.117 \pm 0.111$	$+0.195 \pm 0.147$
	S		-0.417 ± 0.159	-0.210 ± 0.200

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the D^0 recoil mass distribution. According to Monte Carlo studies the fraction of signal events in this sideband is negligible, and we set it to zero. We fit simultaneously the resolution and lifetime parameters, shown in Secs. I and III of Table IV. The fitted values of C_{comb} and S_{comb} are consistent with 0 as expected.

We are now in a position to perform the final Δt fit, in which we release parameters from the signal, continuum and $B\bar{B}$ combinatorial background Δt and resolution models, as detailed in Table IV, always with the convention that parameters free (fixed) in the final Δt fit are shown with (without) a fit uncertainty.

The results are also shown in Fig. 5 and 6 for the kaon (lepton) tagged samples, where we plot the Δt distributions separately for B^0 and \bar{B}^0 tags, together with the time-dependent raw *CP* asymmetry

$$A(\Delta t) = \frac{N_{S_{\text{tag}}=1}(\Delta t) - N_{S_{\text{tag}}=-1}(\Delta t)}{N_{S_{\text{tag}}=-1}(\Delta t) + N_{S_{\text{tag}}=-1}(\Delta t)}.$$
 (22)

For presentation purposes, only data in the restricted signal region $m_{\rm rec} > 1.860 \text{ GeV}/c^2$ are shown in Figs. 5 and 6, in order to amplify signal/background ratio and be able to see the oscillation. The signal fractions in this region become $\approx 24\%$ and $\approx 18\%$ for kaon and lepton tags respectively.



FIG. 5 (color online). Top: Δt distribution for B^0 (dashed) and \bar{B}^0 (solid) kaon tags; the lower curves are the corresponding signal PDF' s. Bottom: raw time-dependent *CP* asymmetry. Only data in the restricted signal region $m_{\rm rec} > 1.860 \text{ GeV}/c^2$ are shown.



FIG. 6 (color online). Top: Δt distribution for B^0 (dashed) and \bar{B}^0 (solid) lepton tags; the lower curves are the corresponding signal PDF' s. Bottom: raw time-dependent *CP* asymmetry. Only data in the restricted signal region $m_{\rm rec} > 1.860 \text{ GeV}/c^2$ are shown.

This requirement is not applied to the fit sample, so our numeric results apply to the whole signal region $m_{\rm rec} > 1.836 \text{ GeV}/c^2$.

V. SYSTEMATIC UNCERTAINTIES

Our systematic uncertainties on the CP-violating parameters S and C are summarized in Table V. We discuss here the most significant ones.

Most systematic uncertainties in Table V are due to imperfect knowledge of one single parameter fixed in the final Δt fit, having little or no correlation with uncertainties of other parameters. They have been treated by varying them by $\pm 1\sigma$ and repeating the final Δt fit leaving only the parameters S and C free to vary.

Uncertainties in the first two lines have a different character because they are due to parameter sets, in which correlations among parameters belonging to one set are nontrivial. Given the low signal-to-background ratio, correct modeling of the background shape and signal fraction in the kinematic fit is crucial, especially because the $\mathcal{M}(m_{\rm rec})$ and $\mathcal{F}(F)$ PDF's parameters are fixed in the final Δt fit. Consequently, we devised a procedure to evaluate the associated systematic uncertainties, that would also preserve the correlations among parameters belonging to a set.

TABLE V. Systematic uncertainties evaluated for C and S. Uncertainties in the top section are independent for kaon and lepton tags, those in the bottom section are correlated.

Systematic source	Kaon tags		Lepto	Lepton tags	
	С	S	C	S	
Kinematic fit parameters	0.013	0.034	0.023	0.057	
Continuum Δt fit parameters	0.002	0.001	• • •	• • •	
Signal <i>s_w</i>	0.0002	0.0007	• • •	• • •	
$B\overline{B}$ combinatorial s_w	0.017	0.0007	0.001	0.005	
Signal tag side (ω)	0.012	0.045	0.002	0.002	
Mistag difference $(\Delta \omega)$	0.007	0.0004	0.007	0.0009	
Signal <i>CP</i> side (α_{D^0})	0.006	0.017	0.002	0.002	
Peaking background	0.0002	0.0003	0.0002	0.00004	
Fit bias (Monte Carlo statistics)	0.011	0.018	0.012	0.019	
Tag interference from DCSD	0.030	0.002			
B^0 lifetime variation	0.0002	0.002	0.0003	0.004	
Δm_d variation	0.0003	0.001	0.0004	0.002	
SVT misalignment	0.003	0.007	0.002	0.004	
Boost uncertainty	0.002	0.006	0.005	0.007	
Total	0.042	0.062	0.028	0.061	

For each set of parameters in \mathcal{M} or \mathcal{F} that become fixed at any stage of our fits, and are not released again in the final Δt fit, a large number N_t of toy Monte Carlo experiments of the same size as the data are generated and fitted, and the values of parameters in the N_t experiments are saved. Evaluation of the systematics due to a set of parameters in subsequent fits (in which they become fixed) is made by repeating the latter fits many times over, using the same event sample, but fixing parameters in the set to different values every time, taken from one of the N_t experiments. In this way, we propagate the variation associated to parameter sets from one fit to the next one, and preserve correct correlations among them. We applied this procedure to obtain the uncertainties in lines 1 and 2 of Table V, which for lepton tagged events are the main source of systematics.

For lepton tags we find that only one Gaussian is sufficient to describe the resolution function $(f^n = 1)$. The systematic due to the signal s^w was evaluated only for kaon-tagged events.

Since the mistag parameter ω is obtained from Monte Carlo with a very small statistical uncertainty (see Table III), we verified the agreement between Monte Carlo and data using a control sample of self-tagging $B^0 \rightarrow D^{*\mp} \pi^{\pm}$ events. As a result of this study, we assign a very conservative uncertainty of 15% on ω , and evaluate the associated systematic by repeating the final Δt fit varying its central value by $\pm 15\%$. This is the largest systematic uncertainty for kaon-tagged events.

We estimate the systematic uncertainty associated with fixing the peaking background fraction f_G in Eq. (8) to zero by setting it to ± 0.002 for both the kaon and lepton tag samples, repeating the fit, and taking the largest deviation from the value fitted with $f_G = 0$ as the systematic uncertainty.

The signal m_{rec} spectra for the *CP*-even and *CP*-odd components are different, the latter being slightly harder. This may cause a small acceptance difference of our event reconstruction and selection, leading to a systematic shift in the *C* and *S* measurement. We have carefully evaluated this effect and found it to be negligible.

As the $D^{*+}D^{*-}$ final state is a superposition of *CP*-even and CP-odd wave functions, the measured values S and Cfrom our data only represent a weighted average of these components, with their inverse squared errors as weights. Since the background shape is not uniform as it goes to zero at the kinematical limit, the weight of the CP-odd component could be enhanced with respect to the CP-even one by the lower background level in the high mass region. To evaluate this effect, we perform Δt fits in the two extreme Monte Carlo configurations, adding to the background sample a pure CP-odd $(R_{\perp} = 1)$ or CP-even $(R_{\perp} = 0)$ sample of signal events, respectively. The number of signal events in both cases is equal to the number of signal events found in data. We find that the differences in the errors of S and C are negligible in these two cases and we do not assign a systematic uncertainty to this effect.

As discussed in Sec. IV C, τ_b and Δm_d are fixed to the values listed in Table III. We assign the systematic uncertainty due to these assumptions by varying their nominal values of $\pm 1\sigma$, and taking half the difference in the fitted values of *C* and *S* so obtained.

To evaluate bias on C and S in our fit, we apply the fit procedure to pure signal Monte Carlo events and compare the results for C and S to the generated ones; no significant bias is observed in either. We therefore quote the statistical uncertainty on these Monte Carlo measurements as the associated systematic uncertainties.

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To measure the systematic uncertainty related to imperfect knowledge of the time measurement due to uncertainty in the boost or possible uncorrected misalignment of the SVT, we repeat the time-dependent fit with different sets of realistic misalignments of the SVT and Δt scaling factors. The maximum observed shift with respect to the nominal fit is quoted as the uncertainty.

An important source of systematic uncertainty in our analysis is represented by interference effects from doubly Cabibbo-suppressed decay (DCSD) amplitudes on the tagging side of the event. The non-leptonic *B*-meson decays used for tagging are dominated by amplitudes containing a $b \rightarrow c\bar{u}d$ transition, thus ensuring the correlation of the tagging particle (typically a kaon) with the flavor of the originating *b* quark. However, $\bar{b} \rightarrow \bar{u}c\bar{d}$ transitions could also contribute, although they are suppressed [22] by a factor $r' \simeq |(V_{ub}^*V_{cd})/(V_{cb}V_{ud}^*)| = 0.02.$

As discussed in detail in Ref. [22], this effect cannot be simply reabsorbed into the mistag probability ω because the allowed and doubly Cabibbo-suppressed amplitudes can interfere, and thus effectively change the Δt probability density function.

Since our Δt PDF assumes r' = 0 and therefore does not include these effects, the *C*, *S* parameters measured by our fit will be different from the observables without tag-side interference by a calculable quantity.

To evaluate the systematic effect in our measurement due to neglecting small terms in the PDF with $r' \neq 0$, we follow the prescription in Ref. [22] and perform a simple toy Monte Carlo of $\delta C \equiv C_{\text{fit}} - C_0$ and $\delta S \equiv S_{\text{fit}} - S_0$, finding the results reported in Table V. The lepton tags are not affected by this issue.

VI. PHYSICS RESULTS

The final results for C and S, with their correlation coefficient ρ , including only the statistical uncertainty for kaon and lepton tags, are:

$$C = +0.12 \pm 0.11$$

 $\rho = 0.0601$, kaon tags,
 $S = -0.42 \pm 0.16$
 $C = +0.20 \pm 0.15$
 $\rho = 0.0730$, lepton tags.
 $S = -0.21 \pm 0.20$

The two samples are statistically independent of each other and can therefore be combined; their statistical uncertainties can be combined in quadrature, but the systematic ones need a more careful treatment.

Indeed, several of the systematic effects listed in Table V are independent for the kaons and lepton tags and are combined in quadrature, while the others are combined taking into account their correlation. Finally we get the combined results of this analysis of

$$C = +0.15 \pm 0.09 \pm 0.04$$

$$\rho = 0.0649.$$

$$S = -0.34 \pm 0.12 \pm 0.05$$

A. Extraction of S_+ and C_+

The measured values of *S* and *C* that we obtain from data only represent a weighted average of the *CP*-even and *CP*-odd wave function components. If penguin amplitudes can be neglected then $S_+ = -S_-$, $C_+ = -C_-$ and the value of the *CP*-even components S_+ and C_+ , which we are interested in, can be obtained using the relations

$$C = C_+ \qquad S = S_+ (1 - 2R_\perp),$$

where the factor $(1 - 2R_{\perp})$ represents the dilution introduced by the *CP*-odd component R_{\perp} in the signal. To compute S_+ we use the value measured by *BABAR* of $(R_{\perp} = 0.158 \pm 0.029)$ [8], where the uncertainty is the combined statistical and systematic. To evaluate the related systematic uncertainty, we vary this value by $\pm 1\sigma$. We obtain

$$C_{+} = +0.15 \pm 0.09 \pm 0.04$$

 $S_{+} = -0.49 \pm 0.18 \pm 0.07 \pm 0.04.$

where the uncertainties shown are statistical and systematic; the third uncertainty is the contribution from the error on R_{\perp} described above.

VII. SUMMARY

We have measured the time-dependent *CP* asymmetry parameters *C* and *S* in $B^0 \rightarrow D^{*+}D^{*-}$ decays, from which we have extracted the *CP*-even components S_+ and C_+ . This result is an independent determination of the *CP*-violating parameters of $b \rightarrow (c\bar{c})d$ transitions and is compatible with previous measurements from *BABAR* [8] and Belle [9] using fully reconstructed decays. It also agrees well with the Standard Model expectation of negligible contributions to the decay amplitude from penguin diagrams and thence with $S_+ = -\sin 2\beta$.

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